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ESTIMATION AND CONFIDENCE INTERVALS FOR
QUANTAL RESPONSE OR SENSITIVITY DATA

by

Guy Burton Seibert, Jr.

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DEPARTMENT OF STATISTICS
Southern Methodist University

ESTIMATION AND CONFIDENCE INTERVALS FOR
QUANTAL RESPONSE OR SENSITIVITY DATA

A Thesis Presented to the Faculty of the Graduate School

of

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in

Partial Fulfillment of the Requirements

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with a

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by

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where $Y_i(\theta_1, \theta_2) = \theta_1 + \theta_2 x_i$, $Y_i = F^{-1}(P_i)$ and $F^{-1}(P_i)$ is the inverse of

$$(3) \quad \sum_{i=1}^t n_i (P_i - P_i(\theta_1, \theta_2))^2 \div \sum_{i=1}^t P_i (P_i(\theta_1, \theta_2) - Y_i)^2$$

estimated. An approximation for the chi-square quantity (1) is given by where x_i is the i th level of stimulus and θ_1, θ_2 are parameters to be

$$(2) \quad P_i(\theta_1, \theta_2) = F(\theta_1 + \theta_2 x_i)$$

response $P_i(\theta_1, \theta_2)$ is given by some distribution function where P_i and $P_i(\theta_1, \theta_2)$ are the observed and expected proportions respectively out of n_i subjects at the i th level of stimulus. The expected

$$(1) \quad \sum_{i=1}^t n_i (P_i - P_i(\theta_1, \theta_2))^2$$

The estimation of the quantal response curve parameters of a quantal bio-assay and the construction of confidence intervals for these parameters are the subjects considered in this paper. The method of minimum chi-square applied to quantal response data requires the minimization of the quantity

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Sensitivity Data

Estimation and Confidence Intervals for Quantal Response or

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problems are worked for both burrit analysis and the maximum likelihood of the Gauss-Newton method for solving a set of non-linear equations. Example where $\theta = (\theta_1, \theta_2, \theta_3 = c, \theta_4 = k)$. The solution to (5) is obtained by use

$$(5) \quad \sum_{i=1}^t \frac{n_i (P_i - P_i(\theta))}{P_i(\theta)} \frac{\partial P_i(\theta)}{\partial \theta_j} = 0, \quad j = 1, 2, 3, 4$$

equations given by

considered in this paper. This requires the solution of a set of four of the Burr distribution by the method of maximum likelihood is also considered. The more general estimation problem of estimating all four parameters by computer.

desk calculator and a FORTRAN IV program is provided for doing the analysis distribution. The appendix contains tables for doing burrit analysis by fidence regions are given for (θ_1, θ_2) and for the LD $_{\alpha}$ point of the Burr tion is (4) instead of normal or logistic. Approximate and exact con- to probit or logit analysis except the form of the underlying distribu- c and k fixed is considered in detail. Burrit analysis is defined parallel The minimization of the approximation (3) using the distribution (4) with

$$(4) \quad F(X) = 1 - (1 + X^c)^{-k}, \quad X > 0, \quad c, k > 1$$

the Burr distribution given by of Y_i on $Y_i(\theta_1, \theta_2)$. The particular choice of $F(Y)$ used in this paper is the estimation problem from a non-linear one to a weighted linear regression the distribution function evaluated at P_i . The approximation (3) changes

solution given by (5). The Burr distribution when estimating all four parameters, offers a greater amount of flexibility in the shape of the quantal response curve than the normal or logistic which only have two parameters to be estimated.

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The function $f(u)$ is a density function to be specified and $Y_i = \theta_i + \alpha_i^2 x_i$.

$$P_i = r_i / n_i \quad .$$

and

$$(1) \quad E(P_i) = P_i = \int_{x_i}^{\infty} f(u) du \quad ,$$

where

$$e^{(r_i)} = \binom{n_i}{r_i} P_i^{r_i} (1 - P_i)^{n_i - r_i} \quad , \quad r_i = 0, 1, \dots, n_i$$

tributton,

n_i when subjected to the dose metameter x_i is given by the binomial distribution, independent of all the others. The probability that r_i respond out of n_i log dose. The reaction of each of the n_i subjects is assumed to be dose levels x_i may be some transformation of the actual dose such as respond, where $i = 1, 2, \dots, t$ and t is the number of dose levels. The subjects are exposed to some level of stimulus x_i and r_i out of the n_i . The general situation for quantal analysis is one in which n_i subjects total number which were subjected to a specified level of stimulus. death. That is, one counts the number of subjects which die out of the characteristic response. A typical type of quantal response often used in situations it is observed whether or not the subject manifests some characteristic of response under the action of a stimulus is not possible. In such In some types of biological experiments a quantitative measurement

INTRODUCTION

$F_1(\theta_1, \theta_2)$ then is the probability that a subject will respond when exposed

to a dose level less than or equal to x_1 . This probability is given by

specifying some absolutely continuous distribution function $F(Y_1)$, so that

$$F_1(\theta_1, \theta_2) = 1 - Q_1(\theta_1, \theta_2) = F(Y_1) = F(\theta_1 + \theta_2 x_1)$$

Throughout the rest of this paper F_1 will be understood to be a function

of θ_1 and θ_2 unless otherwise specified.

The log likelihood function for t doses is given by

$$L = \sum_{t=1}^T \{ \ln c_t + n_t (p_t \ln p_t + q_t \ln Q_t) \} \quad (2)$$

where $c_t = \begin{pmatrix} n_t \\ r_t \end{pmatrix}$. Maximum likelihood estimates for the parameters θ_1, θ_2

are found as the solution to the equations

$$\frac{\partial L}{\partial \theta_j} = \sum_{t=1}^T \frac{p_t Q_t}{n_t (p_t - Q_t)} \frac{\partial \theta_j}{\partial p_t} = 0 \quad (3)$$

These equations are nonlinear in the parameters θ_1, θ_2 and hence cannot

be solved directly. Probit analysis provides for an iterative solution

to (3) where $F(Y)$ is taken as the normal distribution function. A detailed

description of the probit solution is given by Finney [1952]. The solution

makes use of a 'working deviate', Y_1 , defined as

$$Y_1 = \frac{Y_1}{\sigma_1} + \frac{Z_1}{\sigma_1} \quad (4)$$

and weights

$$W_1 = \frac{p_1 Q_1}{n_1 Z_1} \quad (5)$$

where $Y_i = F^{-1}(p_i)$ and $Z_i = \partial F(Y_i) / \partial Y_i$. Each cycle consists of calculating W_i, Y_i from the previous cycle and then forming the weighted linear regression of Y_i on X_i . The maximum likelihood solution (3) when P_i is taken as the logistic distribution has been discussed by Berkson [1955 A]. A linear weighted regression solution has been given by Nelder [1968] which uses the Newton-Raphson iteration method and converges to the maximum likelihood solution.

Another method of estimating a quantal type bio-assay is that of minimum chi-square. Estimates found by this procedure minimize the statistic

$$\chi^2 = \sum_{i=1}^t \frac{e_i}{(o_i - e_i)^2} \quad (6)$$

where o_i and e_i are the observed and expected frequencies of response at the i^{th} dose level. At each dose level the observed and expected frequencies are given by the following table.

	response	non-response	total
observed	p_n	q_n	n
expected	P_n	Q_n	n

The χ^2 at each dose is given by

$$\chi^2 = \frac{p_n}{(p_n - P_n)^2} + \frac{q_n}{(q_n - Q_n)^2}$$

After combining terms and summing over all dose levels we have

$$\chi^2 = \sum_{i=1}^t \frac{P_i Q_i}{n_i (P_i - P_i)^2} \quad (7)$$

As in the case of the maximum likelihood solution (3) the normal equations for (7) are nonlinear in the parameters θ_1, θ_2 . An approximating χ^2 for (7) has been given by Berkson ([1944], [1953], [1955 A]) where P_i is taken to be the logistic distribution,

$$(8) \quad P_i = \frac{1}{1 + \exp[-(\theta_1 + \theta_2 x_i)]}$$

In equation (7), the quantity $(P_i - P_i^*)^2$ is replaced by

$$(9) \quad (P_i - P_i^*)^2 \div (P_i^* Q_i^*) (P_i^* q_i^*) (\lambda_i^* - L_i^*)^2$$

where $\lambda_i^* = \ln(P_i^*/q_i^*)$ and $L_i^* = \ln(P_i^*/Q_i^*) = \theta_1 + \theta_2 x_i^*$ are the observed and estimated logits respectively. Using the approximation (9), equation (7)

may be replaced by a 'minimum logit chi-square',

$$(10) \quad \chi^2 = \sum_{i=1}^t \frac{1}{n_i} (P_i^* - P_i^*)^2 \div \sum_{i=1}^t \frac{1}{n_i P_i^* q_i^*} (\lambda_i^* - L_i^*)^2$$

The minimization of (10) may be viewed as a weighted linear regression of λ_i^* on x_i^* with weights $w_i^* = n_i P_i^* q_i^*$. The minimum χ^2 method has been discussed by Berkson ([1955 B], [1957]) for normal analysis, where P_i^* is taken to be the normal distribution function.

The estimation of θ_1, θ_2 may also be carried out by consideration of the following non-linear regression model,

$$(11) \quad P_i = P_i^* + \epsilon_i$$

with $E(\epsilon_i) = 0$ $\text{Var}(\epsilon_i) = P_i^* Q_i^*/n_i$. If the residual sum of squares $S(\theta)$ is formed using weights $w_i^* = n_i P_i^* Q_i^*$ we have

$$(12) \quad S(\theta) = \sum_{i=1}^t \epsilon_i^2 w_i^* = \sum_{i=1}^t \frac{\epsilon_i^2}{P_i^* Q_i^*}$$

$$(17) \quad \frac{df}{dx} = \frac{b_0 + b_1 x + b_2 x^2}{(x-a)^2}$$

is given in Kendall and Stuart [1963], page 148, as type III, or gamma distribution. The equation for the Pearson system types I, IV, and VI, including many transitional types, such as Burr and Cislak [1968] have shown that this system of distributions have curve shape characteristics which cover a large portion of the Pearson

$$(16) \quad x = \frac{c-1}{ck+1} \left(\frac{1}{c} \right)^{1/c}, \quad c > 1$$

which is unimodal at

$$(15) \quad f(x) = \frac{c^{kx} (1+x)^{c-k-1}}{c^{kx}}$$

with density function,

$$(14) \quad F(x) = 1 - \left[\frac{c}{1+x} \right]^{k-1}, \quad x \geq 0, \quad c, k > 1$$

was originally proposed by Burr [1942] and is given by which will be used to estimate a quantal bio-assay. This distribution In this paper we present a general family of distribution functions which is equivalent to the maximum likelihood normal equations (3).

$$(13) \quad \frac{\partial S(\theta)}{\partial \theta_j} = \sum_{i=1}^t w_i (p_i - P_i) \frac{\partial p_i}{\partial \theta_j} = 0$$

tain for the normal equations is minimized without taking partials with respect to the weights, we obtain which is equivalent to (7). Further (Moore and Zeigler [1967]) if (12)

A generalization of (17) (K. and S., pg. 173) is

$$dF = F(1-F)g(x) dx \quad (18)$$

where $g(x)$ is some convenient function, which must be non-negative within

the range of x and for $0 \leq F \leq 1$. The solution of (18) is

$$F(x) = [1 + \exp\{-G(x)\}]^{-1} \quad (19)$$

where

$$G(x) = \int_x^{-\infty} g(t) dt .$$

If we choose $G(x) = \ln[(1+x)^c]$ then one particular form of (19) is given by the Burr distribution (14).

Two distributions included within the range of approximation by the

Burr are the normal and logistic. Thus, for appropriately chosen values

of c and k both probit and logit analysis may be carried out using the

Burr distribution. The analogous estimation procedure using the Burr dis-

tribution will be termed burrit analysis. The values of c and k which are

used to approximate any given distribution are found by specifying the first

four moments of the given distribution. In the case of the normal distri-

bution these have been given by Burr [1967] as

$$c = 4.874 \quad , \quad k = 6.158 \quad . \quad (20)$$

$$G'(y) = \left\{ n(p-p) \frac{\partial}{\partial p} - \frac{\partial}{\partial p} \left[\frac{\partial}{\partial p} \right] \right\} \Big|_{y=\bar{y}} = 0$$

and

$$G(y) = \frac{n(p-p)^2}{2} = 0$$

The first two terms of this expansion are zero since

$$G(y) = G(y) + (y-\bar{y})G'(y) + \frac{1}{2}(y-\bar{y})^2 G''(y) \quad (22)$$

terms is

The Taylor's expansion of $G(y)$ about the point $y = \bar{y}$ to the first three

$$G(y) = \frac{p\bar{p}}{n} (p-p)^2 \quad (21)$$

We consider first the function defined by

distribution with c and k known.

parameters θ_1, θ_2 and the particular solution will be given for the Burr

function $F(y) = p$. This approximation will be used to estimate the

the (7) which will hold for any choice of absolutely continuous distribution

In this chapter we present an approximation to the chi-square statis-

AND BURRIT ANALYSIS

APPROXIMATE MINIMUM CHI-SQUARE

CHAPTER I

so that (24) is equivalent to Berkson's minimum logit chi-square (10).

$$F'(y) = pq \quad w = npq \quad (26)$$

$$y = \ln(p/q) = L \quad Y = \ln(p/q) = \lambda$$

(8), then we have

is also of interest to note that if p is taken to be the logistic function substituted for theoretical frequencies p , q in equations (4) and (5). It of the maximum likelihood solution if observed frequencies p , q are substituted. These estimates are the same as those resulting from the first iteration

$$w = \frac{np[F'(y)]^2}{pq} \quad (25)$$

linear regression of Y on X with weights

The approximate minimum chi-square (24) may be viewed as a weighted

$$\sum_{i=1}^t \frac{n_i p_i q_i}{n_i [F'(y_i)]^2} \div \sum_{i=1}^t \frac{n_i p_i q_i}{n_i [F'(y_i)]^2} (Y_i - Y)^2 \quad (24)$$

The desired approximation for the minimum chi-square (7) is

$$G(y) = \frac{np}{n[F'(y)]^2} (Y - Y)^2 \quad (23)$$

The approximation for $G(y)$ is now given by

$$G''(y) = \left\{ \frac{\partial^2}{\partial p^2} \cdot \left[\frac{2n}{p} \frac{\partial p}{\partial y} - 2n(p-q) \frac{\partial}{\partial (1/pq)} \right] \right\}_{Y=Y}$$

$$= 2n[F'(y)]^2/pq$$

The third term of (22) is found from the expression for $G''(y)$,

The approximate minimum χ^2 (24) for burrit analysis becomes the minimum

$$\chi^2_{\text{burrit}}$$

$$(32) \quad \sum_{i=1}^t w_i (B_i - b_i)^2$$

where w_i is defined by (31). The estimates for θ_1, θ_2 are given by (28)

with y_i, w_i defined by (29) and (31) respectively. In the actual computing

of (28) it is not necessary to carry along the quantity $c^2 k^2$ in (31) since

this is a constant and will cancel out of the expressions for θ_1, θ_2 . A

table is given in Appendix I which lists the weights and burrits for dif-

ferent values of n_i and responses r_i . This table lists the weights with-

out the quantity $c^2 k^2$ and may be used along with a desk calculator to

compute (28). A Fortran IV computer program which computes (28) and gives

a residual analysis for the least square solution (32) is given in Appendix

II. This program also evaluates the logit analysis solution for θ_1, θ_2

and a modified burrit analysis which will be discussed next.

Burrit Analysis Adjusted for Bias and Variance

As was noted above, (24) may be viewed as a weighted linear regression

of y on x . However, in general it is not possible to define y and weights,

w , such that

$$E(y) = y$$

and

$$w = [\text{var}(y)]^{-1}$$

For logit analysis this problem of bias and variance of the observed logit

has been discussed by Gart and Zweifel [1967].

The mean and variance of y may be given in terms of the Taylor's

The approximate minimum X^2 (24) for burrit analysis becomes the minimum burrit X^2

$$(32) \quad \sum_{i=1}^t w_i (B_i - b_i)^2$$

where w_i is defined by (31). The estimates for θ_1, θ_2 are given by (28)

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Burrit Analysis Adjusted for Bias and Variance

As was noted above, (24) may be viewed as a weighted linear regression

of Y on X . However, in general it is not possible to define Y and weights,

w , such that

$$E(Y) = Y$$

and

$$w = [\text{var}(Y)]^{-1}$$

For logit analysis this problem of bias and variance of the observed logit

has been discussed by Gart and Zweifel [1967].

The mean and variance of Y may be given in terms of the Taylor's

p in (34) and (35) must be replaced by the observed proportion p. The
 with var(Y) defined by (35). To estimate θ_1, θ_2 the expected proportion

$$(37) \quad w^* = [\text{var}(Y)]^{-1}$$

and the weights are given by

$$(36) \quad Y^* = Y - \xi$$

Y on x by regressing Y* on x where

The bias and variance of Y may now be accounted for in the regression of

$$(35) \quad \text{var}(Y) \doteq pQY^{(1)}(p)^2/n + \{2p^2\delta^2/n^2 + pQ(1-6pQ)/n^3\} \\
 \cdot Y^{(2)}(p)^2/4 + pQ\delta^2/n^2 + pQ(1-6pQ)/n^2$$

(33) as

Similarly the var(Y) may be approximated from the first three terms of

$$(34) \quad \xi = E(Y) - Y \doteq pQY^{(2)}(p)/2n + pQ\delta^2/n^2 + pQ(1-6pQ)/24n^3 \\
 + \{p^2\delta^2/8n^2 + pQ(1-6pQ)/24n^3\} Y^{(4)}(p)$$

the binomial an approximation for bias (Y) is

From the first five terms of (33) and the first four central moments of

$$R^{n+1} = \int_p^1 \frac{n!}{(p-t)^n} Y^{(n+1)}(t) dt$$

where

$$(33) \quad Y = Y(p) + \sum_{i=1}^n \frac{Y^{(i)}(p)}{i!} (p-p)^i + R^{n+1}$$

expansion of $Y = F^{-1}(p)$ about the point p

where

$$F_i = F(\theta_1 + \theta_2 x_i)$$

Data Set One			
Burrit	- 0.300915	0.519206	11.756911
Burrit(Adj.)	- 0.296085	0.516622	11.773981
Logit	-10.319148	5.655608	6.340549
	θ_1	θ_2	x^2

Data Set Two			
Burrit	0.518309	0.154945	2.766932
Burrit(Adj.)	0.525622	0.145185	2.997291
Logit	- 1.363051	1.683767	1.735454
	θ_1	θ_2	x^2

Data Set Three			
Burrit	1.124552	0.209810	5.562210
Burrit(Adj.)	1.111075	0.203720	5.585712
Logit	4.841885	2.121552	5.588191
	θ_1	θ_2	x^2

Data Set Four			
Burrit	0.933090	0.151153	1.642129
Burrit(Adj.)	0.927188	0.147964	1.665029
Logit	2.889366	1.519151	1.643307
	θ_1	θ_2	x^2

The above results all give a better fit for the simpler burrit

analysis than for the adjusted burrit analysis. We feel that the additional

amount of approximation required by the adjusted burrit analysis does more harm than good. For this reason it is recommended that burrit analysis be computed in the form given by (32). Also, unless one is sure of the choice for c and k , the values for $\hat{\theta}_1$, $\hat{\theta}_2$ will usually be used as starting values in a non-linear regression to estimate all four parameters of the Burr.

In this chapter we have discussed a general approximation for the χ^2 statistic (7) and have given the specific approximation for the Burr. In Chapter II confidence intervals will be derived for θ_1 , θ_2 and for LD_α point of the quantal response.

$$H_0: \theta_1 = \theta_{10} \quad \text{vs.} \quad H_1: \theta_1 \neq \theta_{10} \quad \theta_2 = \theta_{20} \quad \theta_2 \neq \theta_{20} \quad (43)$$

A test of this hypothesis based on the likelihood ratio statistic is to reject H_0 at the α level of significance if

$$\lambda = \frac{L^*(\theta_1, \theta_2)}{L^*(\theta_{10}, \theta_{20})} < \lambda_0 \quad (44)$$

where

$$P[\lambda < \lambda_0] = \alpha$$

The function L^* is the likelihood function for a quantal bio-assay with binomial response the log of which is given by (2). In the rest of this chapter the notation $\hat{\theta}_1, \hat{\theta}_2$ will refer specifically to the maximum likelihood estimates of θ_1, θ_2 respectively. The maximum likelihood estimates $\hat{\theta}_1, \hat{\theta}_2$ can be found by using the iterative solution given for probit analysis (4) and (5). The burrit analysis estimates can be used as initial values for the estimates $\hat{\theta}_1, \hat{\theta}_2$.

The statistic $-2 \ln \lambda$ will be asymptotically distributed as χ^2 with two degrees of freedom. Using this statistic H_0 is rejected at the α level of significance if

$$-2 \ln \lambda = -2 \ln \frac{L^*(\hat{\theta}_1, \hat{\theta}_2)}{L^*(\theta_{10}, \theta_{20})} > \chi^2_{\alpha} \quad (45)$$

where

$$P[\chi^2_{(2)} > \chi^2_{\alpha}] = \alpha$$

and

$$P_{10} = P(\theta_{10} + \theta_{20} x_1) \quad , \quad P_1 = P(\hat{\theta}_1 + \hat{\theta}_2 x_1)$$

An approximate $100(1-\alpha)\%$ confidence contour for θ_1, θ_2 is the set of all pairs of values $(\theta_{10}, \theta_{20})$ such that (45) is less than or equal to $\chi^2_{2\alpha}$. The above procedure involves iterations to find $\hat{\theta}_1, \hat{\theta}_2$ and also some kind of search procedure for finding the set of $(\theta_{10}, \theta_{20})$ which make up the boundary of the confidence contour. Once the boundary is determined further searching would be necessary to obtain values of $(\theta_{10}, \theta_{20})$ which would give a confidence interval for (41).

The Asymptotic Normal Distribution of $(\hat{\theta}_1, \hat{\theta}_2)$

The second confidence interval will be based on the assumption of a normal distribution for the maximum likelihood estimates $\hat{\theta}_1, \hat{\theta}_2$. Our assumption is that the vector $\hat{\theta}' = (\hat{\theta}_1, \hat{\theta}_2)$ is distributed as

$$\hat{\theta}' \sim \text{BVN}(\bar{\theta}', V) \quad (46)$$

where

$$\bar{\theta}' = (\bar{\theta}_1, \bar{\theta}_2) \quad , \quad V = R^{-1}$$

and

$$R = -E \left[\frac{\partial^2 \log L(\theta_1, \theta_2)}{\partial \theta_i \partial \theta_j} \right]$$

Now from (3) we have that

$$\frac{\partial \log L}{\partial \theta_i} = \sum \frac{\partial \log p}{\partial \theta_i} \quad ,$$

from which

$$\frac{\partial^2 \log L}{\partial \theta_i \partial \theta_j} = \sum \frac{\partial^2 \log p}{\partial \theta_i \partial \theta_j} + \sum \frac{\partial^2 \log p}{\partial \theta_i \partial \theta_j} \frac{\partial \theta_i}{\partial z} \frac{\partial \theta_j}{\partial z}$$

so that

$$R = -E \left(\frac{\partial^2 \log L}{\partial \theta_i \partial \theta_j} \right) = \sum \frac{\partial^2 \log p}{\partial \theta_i \partial \theta_j} \quad (47)$$

$$(52) \quad P \left[(z - Y)^2 / \sigma^2 \leq z^2 \right] = 1 - \alpha .$$

The upper and lower limits on X^α are given by the solution of

$$(53) \quad z^2 - 2zY^\alpha + Y^2 = z^2(V_{11}^\alpha + X^2V_{22}^\alpha + 2X^\alpha V_{12}^\alpha) ,$$

which can be written as a quadratic in X^α

$$(54) \quad X^2 (V_{22}^\alpha - Z^2 V_{22}^\alpha) + 2X^\alpha (\hat{\theta}_1^\alpha - Y^\alpha - V_{12}^\alpha Z^2) + (\hat{\theta}_2^\alpha - Y^\alpha - V_{11}^\alpha Z^2) = 0 .$$

An approximate 100(1- α)% confidence interval for X^α is

$$(55) \quad X^\alpha = \left(Y^\alpha - \hat{\theta}_1^\alpha + V_{12}^\alpha Z^2 - \hat{\theta}_1^\alpha \hat{\theta}_2^\alpha / (\hat{\theta}_2^\alpha - Y^\alpha - V_{11}^\alpha Z^2) \right) \pm \left\{ V_{12}^\alpha Z^2 [V_{12}^\alpha Z^2 + 2\hat{\theta}_2^\alpha (Y^\alpha - \hat{\theta}_1^\alpha)] + V_{22}^\alpha Z^2 [V_{11}^\alpha Z^2 - Y^2 + \hat{\theta}_1^\alpha (2Y^\alpha - \hat{\theta}_1^\alpha)] \right\}^{1/2} / (\hat{\theta}_2^\alpha - Y^\alpha - V_{11}^\alpha Z^2) .$$

The critical point Z^α in (55) may be approximately computed from a t-distribution with $t-2$ degrees of freedom since σ^2 is estimated as a function of $\hat{\theta}_1$ and $\hat{\theta}_2$.

The above confidence interval involves finding $\hat{\theta}_1, \hat{\theta}_2$ by iteration

as did the first interval. Additional computation and approximation is necessary since $V_{11}, V_{12},$ and V_{22} must be estimated. However, a confidence interval for X^α can be found directly without searching the

boundary of the $(\hat{\theta}_1, \hat{\theta}_2)$ confidence contour. The confidence region dis-

cussed next is exact and is based on a non-linear regression model with normal residuals. An approximate confidence region is also given which

requires a minimal amount of computation.

An Exact Confidence Region

The procedure which will be used is based on one given by Hartley

[1964] for constructing exact confidence regions in non-linear estimation.

an exact $100(1-\alpha)\%$ confidence region for θ is

$$Y^t = F(X^t, \theta) + e^t \quad (t = 1, \dots, N) \quad (61)$$

For the non-linear regression model given by

m and $N-m$ degrees of freedom.

where $F(\alpha; m, N-m)$ is the upper $100\alpha\%$ point of the F distribution with

$$\text{reg}(Y-X\theta) / \text{res}(Y-X\theta) < mF(\alpha; m, N-m) / (N-m) \quad (60)$$

An exact $100(1-\alpha)\%$ confidence region for θ is given by

If the residual vector e is distributed $N(0, \sigma^2)$ and the matrix X is of rank m ; then $\text{reg}(e)$ is distributed as χ^2 with m degrees of freedom and $\text{res}(e)$ is independently distributed as χ^2 with $N-m$ degrees of freedom.

$$\text{res}(e) = e'e - \text{reg}(e) \quad (59)$$

and the second component by

$$\text{reg}(e) = (X'e)'(X'X)^{-1}(X'e) \quad (58)$$

The first component of (57) is defined as

$$e'e = \text{reg}(e) + \text{res}(e) \quad (57)$$

error sum of squares $e'e$ into its regression and residual components region for the θ vector may be constructed by a decomposition of the X is a $N \times m$ matrix of known input variables. An ellipsoidal confidence where Y, e are $N \times 1$ matrices, θ is a $m \times 1$ matrix of parameters, and

$$Y = X\theta + e \quad (56)$$

For a linear regression model of the form

For the approximation of $f(x^t, \theta)$ the resulting confidence region may be choice of linear approximation (64). However, depending upon the choice decomposition of $e'e$ will give an exact confidence region for θ for any which is idempotent and of rank m provided U is of rank m . The above

$$(66) \quad A = U(U'U)^{-1}U'$$

The matrix A is now defined to be

$$(65) \quad \begin{aligned} \text{reg}(e) &= (U'e)'(U'U)^{-1}U'e \\ \text{res}(e) &= e'e - \text{reg}(e) \end{aligned}$$

quadratic forms are now given by

$N \times m$ matrix U of rank m which does not depend on θ . The required where the $w_i^t(\theta)$ are a reparameterization of the θ_i and the u_i^t form a

$$(64) \quad f(x^t, \theta) = \sum_{i=1}^m w_i^t(\theta) u_i^t$$

mation of the form

A choice for the matrix A is made by considering a linear approxi-

where A is an idempotent matrix of rank m .

$$(63) \quad \begin{aligned} e'e &= e'[A + (I-A)]e \\ &= \text{reg}(e) + \text{res}(e) \end{aligned}$$

the form

a quadratic form of rank $N-m$. That is, provided the decomposition is of squares $e'e$ in which $\text{reg}(e)$ is a quadratic form of rank m and $\text{res}(e)$ is The confidence region (62) is exact for any decomposition of the sum of

$$(62) \quad \text{reg}(Y-f(X, \theta)) / \text{res}(Y-f(X, \theta)) < \bar{m}F(\alpha; m, N-m) / (N-m)$$

useless. As a measure of the efficiency of the linear model in approxi-
mating the non-linear model we define

$$(67) \quad \frac{\text{reg}(e^*)}{\text{reg}(e)} = \frac{\min Y'Y - e^{*'}e^*}{\min Y'Y - e'e}$$

where e^* is the residual vector associated with the linear model,

$$(68) \quad e^* = Y - \sum_{i=1}^m w_i u_i(\theta)$$

It should be noted that situations can arise in which $\text{reg}(e^*) > \text{reg}(e)$,

e.g., a poor non-linear model or a good non-linear model with high variance
in the data. In such cases (67) may be greater than one, so that in
general there is no upper bound.

The above technique for an exact confidence region will be used for

the quantal response problem by considering the model

$$(69) \quad p_i = p_i + e_i, \quad e_i \sim N(0, \sigma_i^2/n_i)$$

A transformation, which gives a response variable with constant variance,

yields the model

$$(70) \quad z_i = (n_i p_i / \sigma_i^2)^{1/2} + \delta_i, \quad \delta_i \sim N(0, 1)$$

where

$$z_i = p_i / (p_i \sigma_i^2 / n_i)^{1/2}$$

The non-linear function

$$(71) \quad H(y) = (p/\sigma)^{1/2}$$

will be expanded to a first order approximation about the point $y = p^{-1}$.

The $t \times 2$ matrix U containing the transformed input variables u_{1t} , u_{2t} is used to compute

$$z_t - u_{0t} = \theta_1 u_{1t} + \theta_2 u_{2t} + \delta_t^*$$

The above model can now be approximated by

$$z_t - u_{0t} = [(n_t^1 p_t^1 / q_t^1)^{1/2} - u_{0t}] + \delta_t^* \quad , \quad \delta_t^* \sim N(0, 1) \quad (76)$$

the non-linear model (70)

The constant term u_{0t} may be removed by considering a model equivalent to

$$u_{2t} = x_t^1 u_{1t} \quad (75)$$

$$u_{1t} = (n_t^1 p_t^1 / p_t^1)^{1/2} F^1(Y_t^1) / 2 q_t^1$$

$$u_{0t} = (n_t^1 p_t^1 / q_t^1)^{1/2} - Y_t^1 u_{1t}$$

where

$$z_t = u_{0t} + \theta_1 u_{1t} + \theta_2 u_{2t} + \delta_t^* \quad (74)$$

which may be expressed as

$$z_t = (n_t^1 p_t^1 / q_t^1)^{1/2} + (Y_t^1 - Y_t^1) \left(\frac{n_t^1 p_t^1}{p_t^1} \right)^{1/2} \frac{F^1(Y_t^1)}{2 q_t^1} \quad (73)$$

The approximation to (70) is now given by

$$H(Y) = (p/q)^{1/2} \quad , \quad H^1(Y) = \frac{q}{1/2 F^1(Y)} \frac{2 p^{1/2}}{2 q}$$

where

$$H(Y) = H(Y) + (Y - Y) H^1(Y) \quad (72)$$

That is, a linear approximation for $H(Y)$ is

so that an exact $100(1-\alpha)\%$ confidence interval for X^α is

$$(83) \quad X^{\min} < X^\alpha < X^{\max} \quad .$$

The above confidence interval, though exact, requires a computer program which will search out enough values of θ so that the contour

boundary can be graphed and these values used to find X^{\max} , X^{\min} . An

approximation to the above confidence interval and one which involves

much less computing will now be given.

In Chapter I the approximation given by burrit analysis was

$$(84) \quad \sum_{i=1}^t \frac{P_i^2 \theta_i^2}{n_i} (P_i - P_i^2) \div \sum_{i=1}^t \frac{w_i^2 (b_i^2 - B_i^2)}{2} \quad ,$$

where w_i , b_i , and B_i are defined in (29), (30), and (31) respectively.

From (84) we have the linear model

$$(85) \quad b_i = B_i + e_i \quad , \quad e_i \sim N(0, 1/w_i^2) \quad ,$$

in place of the non-linear model (69). This model is equivalent to

$$(86) \quad z_i = B_i w_i^{1/2} + \delta_i \quad , \quad \delta_i \sim N(0, 1) \quad ,$$

where $z_i = b_i w_i^{1/2}$. The model (86) may also be expressed as

$$(87) \quad z_i = \theta u_i^{1/2} + \theta_2 u_i^{2/2} + \delta_i \quad ,$$

where

$$u_i^{1/2} = w_i^{1/2} \quad , \quad u_i^{2/2} = x_i w_i^{1/2} \quad .$$

The $t \times 2$ matrix U of input variables $u_i^{1/2}$, $u_i^{2/2}$ is used to form

$$\text{reg}(\delta) = \delta' U (U' U)^{-1} U' \delta$$

$$(88) \quad = \delta' U' Z - 2 \theta' U' Z + \theta' U' U \theta \quad ,$$

and

$$\text{res}(\delta) = \delta' \delta - \text{reg}(\delta) \tag{89}$$

$$= Z'Z - \hat{\theta}'U'Z, \tag{89}$$

where $\hat{\theta} = (\hat{\theta}_1, \hat{\theta}_2)$ are computed from equation (28).

An approximate $100(1-\alpha)\%$ confidence region for θ is

$$\frac{\text{reg}(\delta)}{\text{res}(\delta)} \leq 2F(\alpha; 2, t-2)/(t-2) = F_\alpha \tag{90}$$

which may also be written as

$$\hat{\theta}'U'U\theta - 2\hat{\theta}'U'Z - \hat{\theta}'Z'Z\hat{\theta} \leq F_\alpha Z'Z - \hat{\theta}'U'Z(F_\alpha+1) \tag{91}$$

The terms needed for the region are

$$\hat{\theta}'U'Z = \hat{\theta}' \int b^I W^I + \hat{\theta}' \int b^I x^I W^I, \tag{92}$$

$$\hat{\theta}'U'Z = \theta' \int b^I W^I + \theta' \int b^I x^I W^I, \tag{92}$$

$$\hat{\theta}'U'U\theta = \theta' \int W^I W^I + 2\theta' \int x^I W^I + \theta' \int x^I x^I W^I, \tag{92}$$

$$Z'Z = \int b^I W^I. \tag{92}$$

The boundary of the confidence region is the set of (θ_1, θ_2) which satisfy

$$s_1^2 \theta_1^2 + s_2^2 \theta_1 \theta_2 + s_3^2 \theta_2^2 + s_4 \theta_1 + s_5 \theta_2 = c \tag{93}$$

where

$$s_1 = \int W^I, \quad s_2 = \int x^I W^I, \quad s_3 = \int x^I x^I W^I, \tag{94}$$

$$s_4 = -2 \int b^I W^I, \quad s_5 = -2 \int b^I x^I W^I, \tag{94}$$

where

$$(99) \quad \frac{A-B}{D} \leq \theta_2 \leq \frac{A+B}{D}$$

found by first picking a θ_2 from the interval
 These conditions imply that values of (θ_1, θ_2) which satisfy (93) may be

$$(98) \quad \theta_2 = \frac{2s_1s_5 - s_2s_4}{2s_1s_5 - 4s_1s_3} \pm \frac{\left[\left(2s_2s_4 - 4s_1s_5 \right)^2 - 4 \left(s_2^2 - 4s_1s_3 \right) \left(s_2^2 + 4s_1c \right) \right]^{1/2}}{2 \left(s_2^2 - 4s_1s_3 \right)}$$

The two values of θ_2 for which equality holds are

$$(97) \quad \left(s_2^2 - 4s_1s_3 \right) \theta_2^2 + (2s_2s_4 - 4s_1s_5) \theta_2 + \left(s_2^2 + 4s_1c \right) \geq 0$$

However, the above solution for θ_1 will be real only if

$$(96) \quad \theta_1 = - \frac{2s_1}{(s_2\theta_2 + s_4)} \pm \frac{\left[(s_2\theta_2 + s_4)^2 - 4s_1(s_3\theta_2^2 + s_5\theta_2 - c) \right]^{1/2}}{2s_1}$$

with solution

$$(95) \quad s_1\theta_1^2 + (s_2\theta_2 + s_4)\theta_1 + (s_3\theta_2^2 + s_5\theta_2 - c) = 0$$

a quadratic in θ_1

interval for X^α . The contour equation (93) can be considered first as

following solution of (93) is for the purpose of constructing a confidence

on the θ_1, θ_2 axis by the usual methods of analytical geometry. The

The confidence contour (93) is the equation of an ellipse and may be plotted

$$(94) \quad c = F \int_0^\alpha b_{1,w_1}^\alpha - (F+1) (\theta_1 \int_0^\alpha d_{1,w_1}^\alpha + \theta_2 \int_0^\alpha d_{1,x,w_1}^\alpha)$$

and

$$A = 4s_1^2 s_5^2 - 2s_2^2 s_4^2 ,$$

$$B = \left[(2s_2^2 s_4^2 - 4s_1^2 s_5^2) - 4 \left(\frac{s_2^2}{s_2^2} - 4s_1^2 s_3^2 \right) \left(\frac{s_4^2}{s_2^2} + 4s_1^2 c \right) \right]^{1/2} .$$

$$D = 2 \left(\frac{s_2^2}{s_2^2} - 4s_1^2 s_3^2 \right) , \tag{100}$$

and then to determine θ_1 from (96).

A confidence interval for X^α may be constructed from (θ_1, θ_2) on the boundary defined by (93). The function $X^\alpha(\theta)$ evaluated along the boundary is defined by

$$X^\alpha(\theta_2) = \frac{2s_1^2 s_5^2 + (s_2^2 + s_4^2) \pm [(s_2^2 + s_4^2) \pm (s_2^2 + s_4^2) - 4s_1^2 (s_2^2 + s_4^2 + s_3^2 - c)]^{1/2}}{2s_1^2 s_2^2} \tag{101}$$

where Y^α is given in (50). An approximate 100(1- α)% confidence interval for X^α is

$$\inf_{\theta_2 \in \Phi} X^\alpha(\theta_2) < X^\alpha < \sup_{\theta_2 \in \Phi} X^\alpha(\theta_2) , \tag{102}$$

where

$$\Phi = \left\{ \theta_2 \mid \frac{A-B}{D} < \theta_2 < \frac{A+B}{D} \right\} . \tag{103}$$

In this chapter we have given several ways of constructing confidence contours for θ and confidence intervals for X^α . The exact confidence region

given by (81) and (83) are felt to be optimum with regard to necessary

assumptions and the exactness of the confidence level. However, these

regions require a program which will search out the boundary values of

θ . If ease of computation is the determining factor in one's choice of a

region, then the approximate regions given by (93) and (102) are recommended.

In estimating all four parameters the flexibility of the Burr is used which offers an advantage over specific choices of distributions such as logistic or normal. By doing this the observations themselves will select the correct distributional form, and it will be unnecessary to do so by assumption. Whenever a fixed choice for c and k is taken, then burrit analysis assumes a fixed distributional form. For the normal approximation of the Burr, burrit analysis is just another way of doing probit or normal analysis. However, even in this case there is an advantage

$$Y = \theta_1 + \theta_2 X$$

and

$$F(Y) = 1 - 1/(1 + Y^{\theta_3})^{\theta_4} \quad (104)$$

given by

$\theta' = (\theta_1, \theta_2, \theta_3, \theta_4)$ will be estimated where the Burr distribution is

four parameters of the Burr distribution. That is, the vector

we give an iterative solution for the maximum likelihood estimates of all

was estimated by burrit analysis holding c and k fixed. In this chapter

In Chapter I, the quantal response curve using the Burr distribution

BURR DISTRIBUTION

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CHAPTER III

in using the Burr in that the distribution function is a closed algebraic

expression and hence also its inverse.

The normal equations for the maximum likelihood solution are

$$\frac{\partial L}{\partial \theta_j} = \sum_{i=1}^n \frac{\partial \ln f_i(\theta_j)}{\partial \theta_j} = \sum_{i=1}^n \frac{\partial \ln f_i(\theta_j)}{\partial \theta_j} \quad (105)$$

$$= \sum_{i=1}^n \frac{\partial \ln f_i(\theta_j)}{\partial \theta_j}, \quad i = 1, 2, 3, 4. \quad (105)$$

These normal equations are equivalent to those for weighted non-linear

least squares if partials are not taken with respect to the weights w_i .

The normal equations are expanded in a first order Taylor's series about

a vector of initial guesses $\theta^0 = (\theta_{10}, \theta_{20}, \theta_{30}, \theta_{40})$. The initial

guesses used for θ_1, θ_2 are those obtained from burrit analysis and $\theta_3,$

θ_4 are taken as values for the normal approximation. Expanding (105) about

θ^0 yields

$$\frac{\partial L}{\partial \theta_j} + \delta \theta_1 \frac{\partial^2 L}{\partial \theta_1 \partial \theta_j} + \delta \theta_2 \frac{\partial^2 L}{\partial \theta_2 \partial \theta_j} + \delta \theta_3 \frac{\partial^2 L}{\partial \theta_3 \partial \theta_j} + \delta \theta_4 \frac{\partial^2 L}{\partial \theta_4 \partial \theta_j} = 0$$

$$i = 1, 2, 3, 4, \quad (106)$$

where $\delta \theta_j = (\theta_j - \theta_j^0)$, $j = 1, 2, 3, 4$ and all partial derivatives

are evaluated at θ^0 . This is a system of four equations which must be

solved for the four $\delta \theta_j$. The normal equations may also be expressed as

$$G_{11} \delta \theta_1 + G_{12} \delta \theta_2 + G_{13} \delta \theta_3 + G_{14} \delta \theta_4 = - \frac{\partial L}{\partial \theta_1} \quad (107)$$

where

$$G_{ij} = \frac{\partial^2 L}{\partial \theta_i \partial \theta_j} = \sum_{i=1}^n \frac{\partial^2 \ln f_i(\theta_j)}{\partial \theta_i \partial \theta_j}$$

$$\begin{aligned}
 P_1 &= CK(V-1)^V \cdot V^{-(K+1)} / Y \\
 P_2 &= X P_1 \\
 P_3 &= K(V-1) \cdot (\ln Y)^V \cdot V^{-(K+1)} \\
 P_4 &= (\ln V)^V \cdot V^{-K} \\
 P_{11} &= CK V^{-(K+2)} (V-1) \{ C(K+1) - V(1+CK) \} / Y^2 \\
 P_{12} &= X P_{11}
 \end{aligned}$$

The necessary first and second partials for the Burr are

which eliminates the need for the second partials P_{1j} .

$$G_{1j}^* = E(G_{1j}) = - \int W P_{1j} \quad (111)$$

Iterations are continued until the process converges in terms of change in both error sum of squares and the δ_j 's. A simplification of (109) may be used provided one is willing to substitute $E(G_{1j})$ for G_{1j} and then solve the resulting normal equations. From (108) the $E(G_{1j})$ is

$$\theta_{j,i} = \theta_{j,i-1} + \delta \theta_j \quad j = 1, 2, 3, 4 \quad i = 2, 3, \dots \quad (110)$$

The quantities G_{1j} are found from (109) and are evaluated at $\theta = \theta_0$. The solutions for $\delta \theta_j$ after the i th iteration give corrections to $\theta_{j,i}$ by

$$P_{1j} = \frac{\partial^2 P}{\partial \theta_j^2} = P_{1j} \quad , \quad P_{1j} = \frac{\partial P}{\partial \theta_j} \quad .$$

and

$$\left[W(P-P) P_{1j} + \left\{ \frac{P Q}{W} \{ P(2P-1) - P^2 \} P_{1j} \right\} \right] = \quad (109)$$

$$\left[W(P-P) P_{1j} + \left\{ \frac{P^2 Q}{(P-P)n(2P-1)} P_{1j} - \int W P_{1j} \right\} \right] = \quad (108)$$

The four sets of data given in Chapter I were used to estimate all four parameters of the Burr. The non-linear program used solved the normal equations (107) and computed the coefficient matrix by (111). The convergence of θ was reached by use of a mixed algorithm which used both the Hartley and Marquardt algorithms. The results of the estimation are given in the following tables which also include the results of burrit analysis for comparison.

$$y = \theta_1 + \theta_2 x^c, \quad \text{and} \quad v = (1 + y^c)$$

where

$$\begin{aligned}
 P_{13} &= P_3 \cdot C[1-K(V-1)]/YV \\
 P_{14} &= P_1 \cdot (1-K \ln V)/K \\
 P_{22} &= X P_{12} \\
 P_{21} &= P_{12} \\
 P_{23} &= X P_{13} \\
 P_{24} &= X P_{14} \\
 P_{31} &= P_{13} + P_1/C \\
 P_{32} &= X P_{31} \\
 P_{34} &= P_3 \cdot (1-K \ln V)/K \\
 P_{33} &= P_3 \cdot (\ln Y) (1-K \cdot Y)/V \\
 P_{41} &= P_{14} \\
 P_{42} &= P_{24} \\
 P_{43} &= P_{34} \\
 P_{44} &= -P_4 \cdot \ln V
 \end{aligned}
 \tag{112}$$

reduction by maximum likelihood does not represent a significantly better

the chi-square for burrit analysis is non-significant, so that a further

logistic nor normal. It should be pointed out that in all of these examples

fulness of the Burr in situations where the response curve is neither

The above group of results, especially data set three, indicate the use-

Max.Lik.	1.167174	0.221299	4.047476	2.631725	1.536194
Logit	2.889366	1.519151	----	----	1.643307
Burrit	0.933090	0.151153	4.874000	6.158000	1.642129
	$\frac{\hat{\theta}_1}{1}$	$\frac{\hat{\theta}_2}{2}$	$\frac{\hat{\theta}_3}{3}$	$\frac{\hat{\theta}_4}{4}$	$\frac{\chi^2}{2}$

Data Set Four

Max.Lik.	1.844753	0.310811	8.668435	0.540272	3.868488
Logit	4.841885	2.121552	----	----	5.588191
Burrit	1.124552	0.209810	4.874000	6.158000	5.562210
	$\frac{\hat{\theta}_1}{1}$	$\frac{\hat{\theta}_2}{2}$	$\frac{\hat{\theta}_3}{3}$	$\frac{\hat{\theta}_4}{4}$	$\frac{\chi^2}{2}$

Data Set Three

Max.Lik.	0.586708	0.420008	3.672498	1.379782	0.778148
Logit	- 1.363051	1.683767	----	----	1.735454
Burrit	0.518309	0.154945	4.874000	6.158000	2.766932
	$\frac{\hat{\theta}_1}{1}$	$\frac{\hat{\theta}_2}{2}$	$\frac{\hat{\theta}_3}{3}$	$\frac{\hat{\theta}_4}{4}$	$\frac{\chi^2}{2}$

Data Set Two

Max.Lik.	- 0.126485	0.565127	7.763276	1.845379	6.808678
Logit	-10.319148	5.655608	----	----	6.340549
Burrit	- 0.300915	0.519206	4.874000	6.158000	11.756911
	$\frac{\hat{\theta}_1}{1}$	$\frac{\hat{\theta}_2}{2}$	$\frac{\hat{\theta}_3}{3}$	$\frac{\hat{\theta}_4}{4}$	$\frac{\chi^2}{2}$

Data Set One

fit. We also note, as in data set one, that there are cases where the Burr does not offer any improvement over the logistic or normal. However, we feel that nothing is lost when using the Burr if all four parameters are estimated and a reduction in the lack of fit χ^2 is likely.

1 - $1/2N$ respectively.

For zero or all responding out of N the response is taken to be $1/2N$ and

$N = 105, 125, 135, \dots, 195$.

$N = 25, 30, 35, \dots, 95$

$N = 6, 7, 8, \dots, 20$

and weights for

tion. If N is the number of subjects, then these tables contain burrits

for burrit analysis when $c = 4.874$ and $k = 6.158$, the normal approxima-

This appendix contains tables of burrits and weights to be used

TABLES OF BURRITS AND WEIGHTS

APPENDIX I

NUMBER	N = 6		N = 7		N = 8		N = 9		N = 10	
RESPONDING	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT
0	.417920	.074388	.404294	.079667	.392923	.084485	.383208	.088936	.374752	.093088
1	.487190	.107571	.470477	.115967	.456661	.123563	.444936	.130536	.434787	.137008
2	.576161	.146785	.553885	.161286	.535926	.174076	.520951	.185613	.508158	.196184
3	.646300	.162790	.617141	.184927	.594535	.203691	.576161	.220178	.560742	.235009
4	.715281	.156548	.675252	.190261	.646300	.217054	.623688	.239692	.605199	.259528
5	.800337	.119402	.736762	.174786	.697423	.213945	.668787	.245065	.646300	.271317
6	.866399	.080112	.816204	.128556	.753968	.190616	.715281	.234823	.686993	.270005
7			.879576	.084556	.829337	.136463	.768253	.204602	.730168	.253515
8					.890652	.088361	.840508	.143411	.780425	.217126
9							.900186	.091679	.850205	.149598
10									.908543	.094614

NUMBER	N = 11		N = 12		N = 13		N = 14		N = 15	
RESPONDING	HURRIT	WEIGHT	HURRIT	WEIGHT	HURRIT	WEIGHT	HURRIT	WEIGHT	HURRIT	WEIGHT
0	•367288	•096989	•360620	•100676	•354606	•104179	•349136	•107520	•344128	•110718
1	•425864	•143067	•417920	•148776	•410775	•154185	•404294	•159334	•398379	•164253
2	•497023	•205981	•487190	•215142	•478404	•223767	•470477	•231935	•463266	•239705
3	•547500	•248568	•535926	•261114	•525668	•272829	•516475	•283849	•508153	•294275
4	•589606	•277328	•576161	•293571	•564368	•308577	•553885	•322571	•544466	•335721
5	•627829	•294276	•612192	•314840	•598664	•333576	•586764	•350862	•576161	•366964
6	•664688	•299668	•646300	•325580	•630486	•348761	•617141	•369854	•605199	•389292
7	•702226	•292531	•680125	•325436	•661854	•354165	•646300	•379844	•659779	•403188
8	•742890	•270447	•715281	•313097	•693390	•349095	•675252	•380522	•659779	•408599
9	•791003	•228459	•753968	•285924	•726675	•332028	•704992	•370987	•686993	•405007
10	•858758	•155167	•800337	•238805	•763758	•300178	•736762	•349573	•715281	•391371
11	•915972	•097240	•866399	•160225	•808677	•248315	•772515	•313386	•780425	•365923
12			•922653	•099610	•873296	•164852	•816204	•257111	•823055	•325689
13					•928718	•101768	•879576	•169112	•885336	•265289
14							•934269	•103746	•885336	•173055
15									•939382	•105569

NUMBER	N = 16		N = 17		N = 18		N = 19		N = 20	
RESPONDING	HURRIT	WEIGHT	HURRIT	WEIGHT	HURRIT	WEIGHT	HURRIT	WEIGHT	HURRIT	WEIGHT
0	•339515	•113789	•335243	•116745	•331268	•119596	•327555	•122355	•324074	•125025
1	•392923	•168969	•387887	•173502	•383208	•177872	•378842	•182092	•374752	•186176
2	•456661	•247126	•450574	•254237	•444936	•261072	•439689	•267657	•434787	•274016
3	•500575	•304188	•493615	•313649	•487190	•322713	•481229	•331419	•475674	•339806
4	•535926	•348151	•528124	•359960	•520951	•371226	•514319	•382011	•508158	•392367
5	•566612	•382078	•557939	•396354	•550001	•409906	•542693	•422826	•535926	•435189
6	•594535	•407381	•584914	•424346	•576161	•440356	•542693	•442826	•535926	•435189
7	•620830	•424681	•610144	•444666	•600487	•463394	•568139	•455545	•560742	•470017
8	•646300	•434107	•634373	•457577	•623688	•479384	•614020	•481057	•583614	•497803
9	•671612	•435395	•658192	•462992	•646300	•488371	•635632	•499807	•605199	•519056
10	•697423	•427889	•682139	•460511	•668787	•490129	•656940	•517358	•625968	•533999
11	•724510	•410450	•706790	•449380	•691607	•484163	•700198	•515742	•646300	•542634
12	•753968	•381232	•738866	•395625	•715281	•469645	•723037	•506524	•666531	•544768
13	•787630	•337200	•767630	•348014	•740491	•445309	•747495	•461330	•708053	•507030
14	•829337	•272925	•794239	•348014	•800337	•409204	•774566	•422055	•730168	•476540
15	•890652	•178722	•835132	•280085	•840508	•286821	•805995	•367844	•780425	•434251
16	•944120	•107257	•895586	•180148	•900186	•183359	•904494	•293179	•811269	•376982
17			•948533	•108828	•952661	•110296	•956537	•111672	•850205	•299196
18									•908543	•189228
19									•960190	•112967
20										

NUMBER	N = 25		N = 30		N = 35		N = 40		N = 45	
RESPONDING	HURRIT	WEIGHT	HURRIT	WEIGHT	HURRIT	WEIGHT	HURRIT	WEIGHT	HURRIT	WEIGHT
0	.309383	.137297	.297904	.148165	.288546	.157995	.280688	.167014	.273940	.175384
1	.357538	.204899	.344128	.221437	.333220	.236366	.324074	.250051	.316229	.262734
2	.414257	.303030	.398370	.328506	.385505	.351414	.374752	.372351	.365552	.391718
3	.452550	.378449	.434787	.411025	.420471	.440724	.440854	.467785	.398370	.492759
4	.482679	.439032	.463266	.479409	.447703	.515375	.454878	.548033	.423794	.578098
5	.508158	.490459	.487190	.537854	.470477	.579837	.456661	.617814	.444936	.652680
6	.530650	.534132	.508158	.588551	.490341	.636456	.475674	.679611	.463265	.719114
7	.551095	.571215	.527064	.632802	.508158	.686642	.492665	.734924	.479603	.778976
8	.570082	.602425	.544466	.671442	.524465	.731313	.508158	.784734	.494454	.833305
9	.588017	.628205	.560742	.705025	.539623	.771095	.522436	.829723	.508158	.882825
10	.605199	.648820	.576161	.733927	.553885	.806428	.535926	.870378	.520921	.928054
11	.621860	.664391	.590926	.758395	.567442	.837630	.548629	.907060	.533010	.969411
12	.638198	.674931	.605199	.778584	.580438	.864922	.560742	.940034	.544466	.1.007123
13	.654387	.680356	.619111	.794573	.592992	.888462	.572374	.969500	.555423	.1.041557
14	.670594	.680479	.632777	.806375	.605199	.924636	.583614	.995605	.565966	.1.072745
15	.686993	.675011	.646300	.813951	.617141	.946432	.594535	.1.018453	.586056	.1.100891
16	.703773	.663540	.659779	.817198	.628891	.937338	.605199	.1.038112	.595731	.1.126021
17	.721161	.645500	.673309	.815959	.640516	.946432	.615660	.1.054623	.595731	.1.148353
18	.739442	.6220128	.686993	.810014	.652077	.951860	.625968	.1.067999	.605199	.1.167876
19	.759006	.586384	.700941	.799068	.663635	.953528	.636167	.1.078225	.614506	.1.184562
20	.740425	.542814	.715281	.782742	.675252	.951304	.646300	.1.085268	.623688	.1.1.198460
21	.804619	.487312	.730168	.730168	.686993	.945016	.656408	.1.089066	.632777	.1.1.209564
22	.833250	.416022	.745797	.731846	.698928	.934443	.666531	.1.089535	.641802	.1.1.217851
23	.869932	.325176	.762428	.695822	.711136	.919312	.676712	.1.086565	.650793	.1.1.223281
24	.925756	.201428	.780425	.651377	.723710	.899283	.686993	.1.080018	.659779	.1.1.225796
25	.975820	.118465	.800337	.597011	.736762	.873932	.697423	.1.069724	.668787	.1.1.225324
26			.823055	.530577	.750433	.842733	.708053	.1.055474	.666923	.1.1.221770
27			.850205	.448794	.764905	.805021	.718945	.1.037020	.666923	.1.1.215021
28			.885336	.346109	.780425	.759940	.730168	.1.014060	.666255	.1.1.204941
29			.939382	.211137	.797347	.706367	.741808	.986229	.705671	.1.1.191359
30			.988302	.122791	.816204	.642778	.753968	.953081	.715281	.1.1.174113
31					.837869	.567007	.766782	.914067	.735284	.1.1.152948
32					.863942	.475771	.780425	.868503	.735284	.1.1.127606
33					.897925	.363557	.795139	.815515	.745797	.1.1.097769
34					.950630	.219149	.811269	.753964	.756753	.1.1.063057
35					.998672	.126327	.829337	.682313	.768253	.1.1.023010
36							.850205	.598392	.780425	.977066
37							.875452	.498929	.793442	.924527
38							.908543	.378456	.807539	.864509
39							.960190	.225933	.823055	.795866
40							1.007531	.129298	.840508	.717054
41									.860747	.625905
42									.885336	.519164
43									.917706	.391417
44									.968491	.231792
45									1.015255	.131847

NUMBER	RESPONDING	N = 50			N = 55			N = 60			N = 65			N = 70		
		BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	
0	HURRIT	•183215	•262827	•190593	•258153	•197555	•253028	•204236	•250080	•210566						
1	•309383	•274594	•303325	•285759	•297904	•296330	•293006	•306385	•288546	•315990						
2	•357538	•409798	•350458	•426797	•344128	•442873	•338417	•458153	•472733	•472733						
3	•399524	•510031	•381719	•537883	•374752	•558527	•368472	•578126	•362764	•596815						
4	•414257	•506061	•405858	•632276	•398370	•657012	•391628	•680475	•385505	•702827						
5	•434787	•685041	•425864	•715334	•417920	•743879	•410775	•770924	•404294	•796668						
6	•452550	•755697	•443146	•789881	•434787	•822049	•427277	•852496	•420471	•881449						
7	•468349	•819675	•458492	•857635	•449743	•893304	•441893	•927025	•434787	•959057						
8	•482679	•878064	•472387	•919728	•463266	•958818	•455094	•995725	•447703	•1.030749						
9	•495872	•931626	•485155	•976958	•475674	•1.019417	•467191	•1.059453	•459528	•1.097404						
10	•508158	•980918	•497023	•1.029904	•487190	•1.075708	•478404	•1.118835	•459528	•1.097404						
11	•519708	•1.026357	•508158	•1.079010	•497977	•1.128150	•488894	•1.174351	•470477	•1.159673						
12	•530650	•1.058264	•518684	•1.124614	•508158	•1.177101	•498780	•1.226373	•480768	•1.218046						
13	•541086	•1.106889	•528701	•1.166985	•517828	•1.222844	•508158	•1.275193	•499466	•1.324575						
14	•551095	•1.142430	•538285	•1.206335	•527064	•1.265605	•517100	•1.321053	•508158	•1.373285						
15	•560742	•1.175042	•547500	•1.242838	•535926	•1.305568	•525668	•1.364147	•516475	•1.419246						
16	•570082	•1.204850	•556398	•1.276630	•544466	•1.342884	•533911	•1.404634	•524465	•1.462627						
17	•579161	•1.231949	•565022	•1.307826	•552726	•1.377678	•541870	•1.442649	•532170	•1.503568						
18	•588017	•1.256413	•573411	•1.336515	•560742	•1.410051	•549580	•1.478303	•539623	•1.542190						
19	•596686	•1.278297	•581596	•1.362768	•568542	•1.440087	•557071	•1.511689	•546853	•1.578591						
20	•605199	•1.297640	•589606	•1.386641	•576161	•1.467854	•564368	•1.542883	•553885	•1.612857						
21	•613582	•1.314465	•597466	•1.408177	•583614	•1.493408	•571495	•1.571952	•560742	•1.645059						
22	•621860	•1.328781	•605199	•1.427404	•590125	•1.516791	•578470	•1.578259	•574742	•1.675259						
23	•630058	•1.340586	•612824	•1.444343	•598115	•1.538036	•585314	•1.623911	•574002	•1.703508						
24	•638198	•1.349863	•620362	•1.459400	•605199	•1.557168	•592040	•1.646881	•580438	•1.729845						
25	•646300	•1.356985	•627829	•1.471378	•612192	•1.574222	•598664	•1.667881	•586764	•1.754308						
26	•654387	•1.360112	•635243	•1.481462	•619111	•1.589146	•605199	•1.686932	•592992	•1.776924						
27	•662477	•1.362192	•642621	•1.489233	•625968	•1.601998	•611657	•1.704047	•599133	•1.797715						
28	•670594	•1.360959	•649977	•1.494664	•632777	•1.612752	•618051	•1.719231	•605199	•1.816697						
29	•678758	•1.356934	•657327	•1.497715	•639550	•1.621394	•624390	•1.732486	•611198	•1.833880						
30	•686993	•1.350023	•664688	•1.498340	•646300	•1.627992	•630686	•1.743806	•617141	•1.849271						
31	•695322	•1.340114	•672074	•1.496481	•653039	•1.632248	•636949	•1.753182	•623036	•1.862871						
32	•703773	•1.327079	•679504	•1.492069	•659779	•1.634395	•643187	•1.760596	•628891	•1.874675						
33	•712375	•1.310766	•686993	•1.485025	•666531	•1.634303	•649411	•1.766026	•634715	•1.884677						
34	•721161	•1.290999	•694560	•1.475257	•673309	•1.631918	•655631	•1.769447	•640516	•1.892864						
35	•730168	•1.267575	•702226	•1.462557	•680125	•1.627182	•661854	•1.770825	•646300	•1.899219						
36	•739442	•1.240256	•710012	•1.447105	•686993	•1.620027	•668093	•1.770121	•652077	•1.903720						
37	•749033	•1.208763	•717942	•1.428453	•693926	•1.610375	•674435	•1.767289	•657852	•1.906343						
38	•759006	•1.172767	•726044	•1.406559	•700994	•1.598137	•680693	•1.762279	•657852	•1.905825						
39	•769437	•1.131879	•734348	•1.381222	•708053	•1.583211	•686993	•1.755030	•669432	•1.905825						
40	•780425	•1.085628	•742890	•1.352233	•715281	•1.565484	•693390	•1.745476	•675252	•1.902609						
41	•792096	•1.033445	•751712	•1.319348	•722666	•1.544846	•699856	•1.733541	•681103	•1.897362						
42	•804619	•.974624	•760866	•1.282281	•730186	•1.521059	•706402	•1.719141	•686993	•1.890032						
43	•818225	•.908274	•770812	•1.242281	•737816	•1.494109	•713044	•1.702181	•692931	•1.880562						
44	•833250	•.833244	•780425	•1.194191	•745797	•1.463692	•719796	•1.682555	•698828	•1.868887						
45	•850205	•.747990	•791003	•1.142297	•753968	•1.429621	•726675	•1.660141	•704992	•1.854934						
46	•869932	•.650351	•802269	•1.084435	•762428	•1.391663	•733170	•1.634806	•711136	•1.838625						
47	•893980	•.537093	•814390	•1.019894	•771227	•1.349468	•740895	•1.606397	•717371	•1.819868						
48	•925756	•.402855	•827595	•.947781	•780425	•1.302754	•748283	•1.574743	•723710	•1.798565						

49	•975820	•236429	•842216	•866942	•790097	1.251099	•755893	1.539648	•730168	1.774605
50	•1.022099	•134068	•858758	•775837	•800337	1.124023	•763758	1.500889	•736762	1.747864
51			•878058	•672312	•811269	1.130946	•771920	1.458211	•743510	1.718203
52			•901853	•553157	•823055	1.061154	•780425	1.411317	•750433	1.685467
53			•932925	•413068	•835924	•983753	•789334	1.359863	•757555	1.649479
54			•982376	•241490	•850205	•897588	•798720	1.303443	•764905	1.610042
55			•1.028239	•136030	•866399	•801124	•808677	1.241575	•772515	1.566928
56					•885336	•692219	•819325	1.173676	•780425	1.519880
57					•908543	•567685	•830829	1.099031	•788683	1.468598
58					•939382	•422275	•843412	1.016744	•797347	1.412734
59					•988302	•245581	•857401	•925663	•806490	1.351882
60					1.033803	•137782	•873296	•824259	•816204	1.285555
61							•891919	•710400	•826610	1.213171
62							•914790	•580926	•837869	1.134013
63							•945252	•249282	•850205	1.047187
64							•993707	•139359	•863942	•951543
65							1.038890		•879576	•845558
66									•897925	•727113
67									•920500	•593074
68									•950630	•438298
69									•998672	•252654
70									1.043572	•140790

RESPONDING	NUMBER	N = 75		N = 80		N = 85		N = 90		N = 95	
		BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT
0		•24551	•216678	•243296	•222527	•240278	•228167	•237468	•234616	•234887	•238887
1		•284458	•325188	•280688	•334029	•277194	•342547	•273940	•350768	•270899	•358725
2		•328460	•486692	•324074	•500101	•320011	•513008	•316229	•525468	•312696	•537515
3		•357538	•614696	•352726	•631859	•348272	•648378	•344128	•664310	•340259	•679710
4		•379906	•724201	•374752	•744702	•369985	•764423	•365552	•783436	•361415	•801810
5		•398370	•821265	•392923	•844845	•387887	•867512	•383208	•889360	•378842	•910458
6		•414257	•909091	•408547	•935570	•403270	•961011	•398370	•985518	•393801	•1.009177
7		•428304	•989616	•422351	•1.018869	•416854	•1.046960	•411753	•1.074004	•406998	•1.100099
8		•440967	•1.084133	•434787	•1.096066	•429083	•1.126710	•423794	•1.156195	•418866	•1.184634
9		•452550	•1.133546	•446153	•1.168091	•440254	•1.201217	•434787	•1.233074	•429696	•1.263784
10		•463266	•1.198523	•456661	•1.235628	•450574	•1.271185	•444936	•1.305359	•439689	•1.338286
11		•473270	•1.259573	•466463	•1.299202	•460194	•1.337151	•454392	•1.373600	•448995	•1.408701
12		•482679	•1.317096	•475674	•1.359222	•469229	•1.399535	•463266	•1.438227	•457724	•1.475467
13		•491585	•1.371406	•484385	•1.416015	•477766	•1.458669	•471647	•1.499584	•465963	•1.538938
14		•500058	•1.422763	•492666	•1.469849	•485876	•1.514832	•479603	•1.557952	•473779	•1.599401
15		•508158	•1.471377	•500575	•1.520938	•493615	•1.568247	•487190	•1.613563	•481229	•1.657097
16		•515931	•1.517426	•508158	•1.569469	•501030	•1.619104	•494454	•1.666610	•495202	•1.712221
17		•523417	•1.561057	•515454	•1.615594	•508158	•1.667560	•501433	•1.717261	•495202	•1.764944
18		•530650	•1.602396	•522496	•1.659445	•515032	•1.713755	•508158	•1.765652	•501793	•1.815411
19		•537658	•1.641548	•529312	•1.701135	•521680	•1.757802	•514656	•1.811908	•508158	•1.863744
20		•544466	•1.678605	•535926	•1.740757	•528124	•1.799801	•520951	•1.856128	•514319	•1.910053
21		•551095	•1.713645	•542359	•1.778394	•534387	•1.839801	•527064	•1.898407	•520298	•1.954432
22		•557563	•1.746734	•548629	•1.814119	•540485	•1.877997	•533010	•1.938821	•526110	•1.996960
23		•563887	•1.77927	•554752	•1.847993	•546433	•1.914333	•544666	•1.977441	•531771	•2.037713
24		•570082	•1.807276	•560742	•1.888068	•552247	•1.948901	•544066	•2.014327	•537295	•2.076751
25		•576161	•1.834818	•566612	•1.910391	•557939	•1.981770	•550001	•2.049531	•542693	•2.114132
26		•582136	•1.860590	•572374	•1.939001	•563519	•2.012963	•555423	•2.083100	•547976	•2.149907
27		•588017	•1.884619	•578039	•1.965932	•568997	•2.042525	•560742	•2.115076	•553154	•2.184115
28		•593815	•1.906930	•583614	•1.991210	•574384	•2.070489	•565966	•2.145495	•558235	•2.216800
29		•599540	•1.927539	•589110	•2.014862	•579687	•2.096883	•571103	•2.1174388	•563228	•2.247993
30		•605199	•1.946461	•594535	•2.036906	•584914	•2.121728	•576161	•2.1201782	•568139	•2.277725
31		•610800	•1.963704	•599895	•2.057356	•590073	•2.145048	•581147	•2.227700	•572975	•2.306023
32		•616351	•1.979274	•605199	•2.076225	•595169	•2.166856	•586066	•2.252162	•577743	•2.332909
33		•621860	•1.993172	•610451	•2.093521	•600209	•2.187167	•590926	•2.275186	•582447	•2.358404
34		•627333	•2.005396	•615660	•2.109247	•605199	•2.205989	•595731	•2.2296786	•587094	•2.382524
35		•632777	•2.015940	•620830	•2.123407	•610144	•2.223331	•600487	•2.316972	•591688	•2.405285
36		•638198	•2.024794	•625968	•2.135998	•615049	•2.239196	•605199	•2.335753	•596234	•2.4426698
37		•643602	•2.031945	•631078	•2.147015	•619920	•2.253986	•609870	•2.353136	•600736	•2.466772
38		•648997	•2.037377	•636167	•2.156451	•624762	•2.266499	•614506	•2.369123	•605199	•2.465517
39		•654387	•2.041068	•641240	•2.164296	•629578	•2.277935	•619111	•2.383719	•609625	•2.489034
40		•659779	•2.042994	•646300	•2.170536	•6349153	•2.287883	•628242	•2.408725	•618386	•2.513813
41		•665179	•2.043124	•651355	•2.175154	•639153	•2.296338	•632797	•2.419128	•622727	•2.539406
42		•670594	•2.041434	•656408	•2.178132	•6436920	•2.308719	•637276	•2.428124	•627046	•2.5539406
43		•676030	•2.037987	•661465	•2.179446	•648679	•2.308719	•641802	•2.435703	•631347	•2.550217
44		•681494	•2.032434	•666531	•2.179070	•653435	•2.312616	•641802	•2.441853	•635633	•2.559695
45		•686993	•2.025034	•671612	•2.176976	•658192	•2.314961	•650793	•2.444853	•639906	•2.567834
46		•692534	•2.015637	•676712	•2.173131	•662954	•2.315731	•655285	•2.446562	•644171	•2.574625
47		•698124	•2.004186	•681837	•2.167498	•667725	•2.314904	•659779	•2.449814	•648429	•2.580056
48		•703773	•1.990619	•686993	•2.160036	•672510	•2.312453				

49	•7094489	1.974867	•692186	2.150702	•677313	2.308347	•664278	2.451878	•652685	2.584116
50	•715281	1.956855	•697423	2.139447	•682139	2.302554	•668787	2.450647	•656940	2.586789
51	•721161	1.936499	•702709	2.126215	•686993	2.295039	•673309	2.447877	•661199	2.588059
52	•727139	1.913705	•708053	2.110948	•691879	2.285761	•677848	2.4443539	•665464	2.587908
53	•733228	1.888372	•713462	2.093581	•696804	2.274677	•682408	2.437605	•669738	2.586315
54	•739442	1.860384	•718945	2.074440	•701772	2.261740	•686993	2.430041	•674025	2.583258
55	•745797	1.829615	•724510	2.052249	•706790	2.246898	•691607	2.420813	•678327	2.578712
56	•752312	1.795924	•730168	2.028120	•711864	2.230093	•696255	2.409882	•682649	2.572650
57	•759006	1.759151	•735930	2.001558	•717001	2.211264	•700941	2.397205	•686993	2.565043
58	•765904	1.719119	•741808	1.972458	•722208	2.190344	•705671	2.382738	•691363	2.555860
59	•773033	1.675626	•747815	1.940702	•727494	2.167258	•710449	2.366429	•695764	2.545065
60	•780425	1.628443	•753968	1.906161	•732866	2.141926	•715281	2.348226	•700198	2.532620
61	•788121	1.577309	•760283	1.868692	•738335	2.114257	•720174	2.328070	•704671	2.518487
62	•796166	1.521923	•766782	1.828134	•743912	2.084156	•725134	2.305896	•709186	2.502619
63	•804619	1.461936	•773487	1.784306	•749609	2.051515	•730168	2.281635	•713749	2.484971
64	•813551	1.376938	•780425	1.737005	•755438	2.016215	•735284	2.255212	•718364	2.465490
65	•823055	1.329443	•787630	1.686001	•761416	1.978125	•740491	2.226543	•723037	2.444421
66	•833250	1.249866	•795139	1.631029	•767559	1.937098	•745797	2.195538	•727774	2.420803
67	•844296	1.166492	•803000	1.5677928	•773888	1.892974	•751214	2.162098	•732581	2.395471
68	•856416	1.075427	•811269	1.497928	•780425	1.845568	•756753	2.126114	•737465	2.368054
69	•869932	•975527	•820017	1.439036	•787198	1.794677	•762428	2.087465	•742434	2.338475
70	•885336	•865273	•829337	1.384626	•794239	1.740069	•768253	2.046019	•747495	2.306649
71	•903442	•742563	•839346	1.284114	•801585	1.681478	•774245	2.001629	•752659	2.272485
72	•925756	•604243	•852005	1.196785	•809283	1.618602	•780425	1.954131	•757936	2.235881
73	•955590	•445346	•862134	1.101746	•817389	1.551088	•786815	1.903341	•763337	2.196728
74	1.003263	•255746	•875452	•997858	•825976	1.478525	•793442	1.849053	•768875	2.154900
75	1.047909	•142098	•890652	•883611	•835132	1.400426	•800337	1.791034	•774566	2.110276
76			•908543	•756913	•844977	1.316206	•807539	1.729019	•780425	2.062694
77			•930622	•614678	•855669	1.225151	•815093	1.662702	•786473	2.011994
78			•960190	•451866	•867426	1.126374	•823055	1.591732	•792732	1.957991
79			1.007531	•258596	•880569	1.018737	•831497	1.515696	•799229	1.900479
80			1.051947	•143300	•895586	•900738	•840508	1.434107	•805995	1.839222
81					•913282	•770300	•850205	1.346383	•813068	1.773956
82					•935150	•624359	•860747	1.251810	•820493	1.704374
83					•964477	•457927	•872351	1.149502	•828326	1.630123
84					1.011517	•261237	•900186	1.038328	•836637	1.550793
85					1.055724	•144409	•917706	•916794	•845517	1.465895
86							•939382	•782834	•855081	1.374848
87							•968491	•633412	•865488	1.276942
88							•98491	•463583	•876953	1.171292
89							1.015256	•263694	•889794	1.056771
90							1.059271	•145438	•904494	•931895
91									•921854	•794612
92									•943354	•641906
93									•972263	•468881
94									1.018775	•265987
95									1.062615	•146395

NUMBER	N=105		N=115		N=125		N=135		N=145	
	RESPONDING	HURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT
0		230052	•248950	•225787	•258460	•221948	•267493	•218463	•276116	•283362
1		265361	•373915	•260428	•388265	•255989	•401898	•251961	•414891	•248279
2		306267	•300512	•300543	•582231	•295955	•602838	•290726	•622482	•286459
3		333220	•709099	•326959	•736832	•321332	•763138	•316222	•788202	•311569
4		353894	•836843	•347208	•869883	•341202	•901214	•335760	•931051	•330790
5		370910	•950671	•363864	•988576	•357586	1.024494	•351808	1.058688	•345779
6		385505	1.054241	•378142	1.096690	•371534	1.138898	•365552	1.175154	•360095
7		398370	1.149771	•390718	1.189635	•383856	1.240802	•377646	1.282903	•371984
8		409931	1.238737	•402012	1.289630	•394915	1.337783	•388496	1.383562	•382645
9		420471	1.322173	•412302	1.377060	•404985	1.428961	•398370	1.478277	•392345
10		430189	1.400847	•421230	1.459613	•414257	1.515152	•407459	1.567897	•401268
11		439230	1.475345	•430595	1.537899	•422872	1.595981	•415899	1.653067	•409551
12		447703	1.546124	•438849	1.612395	•430936	1.674947	•423794	1.734293	•417296
13		455693	1.613553	•446627	1.683480	•438529	1.744436	•431225	1.811985	•424583
14		463266	1.677932	•453993	1.751466	•445716	1.824780	•438255	1.886476	•431474
15		470477	1.739510	•461002	1.816611	•452950	1.882423	•444936	1.958039	•438019
16		477370	1.798492	•467696	1.879136	•459073	1.955048	•451309	2.026913	•444260
17		483981	1.855062	•474111	1.939223	•465320	2.018389	•457410	2.093287	•450231
18		490341	1.909367	•480277	1.997031	•471320	2.079429	•463266	2.157341	•455961
19		496476	1.961539	•486220	2.052696	•477100	2.138311	•468904	2.219216	•461473
20		502408	2.011693	•491962	2.106339	•482679	2.195161	•474343	2.279039	•466790
21		508158	2.059927	•497821	2.158062	•488078	2.250083	•479603	2.336928	•471928
22		513741	2.106331	•502915	2.207960	•493312	2.303177	•486699	2.392978	•476929
23		519172	2.150979	•508158	2.256111	•498395	2.355227	•488646	2.447281	•481731
24		524465	2.193940	•513362	2.302591	•503340	2.404212	•494454	2.499915	•486421
25		529631	2.235277	•518538	2.347464	•508158	2.452295	•499136	2.550951	•490985
26		534661	2.275042	•523097	2.390786	•512858	2.498840	•503701	2.600454	•495433
27		539623	2.313285	•527848	2.432610	•517450	2.543902	•499158	2.648478	•499773
28		544466	2.350048	•532499	2.472844	•521941	2.587529	•512514	2.695080	•504012
29		549218	2.385369	•537057	2.511949	•526339	2.629769	•516776	2.740304	•508158
30		553885	2.419285	•541529	2.549543	•530650	2.670659	•520951	2.784193	•512216
31		558475	2.451826	•545922	2.585802	•534880	2.710239	•525045	2.826787	•516194
32		562992	2.483019	•550239	2.624753	•539039	2.746541	•529063	2.868122	•520094
33		567442	2.512889	•554488	2.664427	•543120	2.785596	•533010	2.908232	•523924
34		571830	2.541458	•558673	2.686850	•547138	2.821431	•538890	2.947145	•527686
35		576161	2.568745	•562797	2.718042	•551095	2.850767	•544707	2.984890	•531385
36		580438	2.594767	•566865	2.748026	•554894	2.889547	•544666	3.021490	•535025
37		584667	2.619540	•570881	2.778811	•558839	2.921871	•548169	3.056970	•538609
38		588850	2.643076	•574849	2.804440	•562633	2.953062	•551821	3.091351	•542140
39		592992	2.665386	•578771	2.830700	•566380	2.983126	•555423	3.124650	•545621
40		597095	2.686480	•582451	2.856215	•570082	3.012142	•559280	3.156889	•549055
41		601163	2.706364	•586491	2.880393	•573742	3.044025	•562493	3.188082	•552445
42		605199	2.725045	•590295	2.903034	•577364	3.068855	•565966	3.218243	•555793
43		609205	2.742527	•594066	2.925388	•580948	3.092688	•569399	3.247388	•559102
44		613185	2.758814	•597605	2.946218	•584499	3.117349	•572797	3.275527	•562373
45		617141	2.773907	•601515	2.965974	•588017	3.141032	•576161	3.302672	•565608
46		621076	2.787805	•605199	2.984573	•591506	3.163864	•579492	3.328834	•568810
47		624992	2.800508	•608858	3.002108	•594966	3.185312	•582794	3.354023	•571980
48		628891	2.812013	•612494	3.018550	•598400	3.205922	•586066	3.378243	•575120

49	•632777	2.4822316	•616111	3.033902	•601811	3.225517	•589313	3.401505	•578222	3.564879
50	•636651	2.8331413	•619709	3.048164	•605199	3.244101	•592534	3.423815	•581317	3.590469
51	•640516	2.839296	•623291	3.061335	•608566	3.261680	•595731	3.445179	•584377	3.615157
52	•644373	2.845958	•626859	3.073414	•611914	3.278253	•598907	3.465600	•587413	3.638946
53	•648226	2.851390	•630413	3.084396	•615245	3.293823	•602062	3.485022	•590426	3.661848
54	•652077	2.855581	•633957	3.094281	•618560	3.308390	•605199	3.503629	•593418	3.683863
55	•655927	2.858519	•637492	3.103061	•621462	3.313263	•608317	3.521244	•596390	3.704998
56	•659779	2.860192	•641019	3.110723	•625148	3.334513	•611419	3.537929	•599244	3.725257
57	•663635	2.8660584	•644541	3.117288	•628424	3.346066	•614906	3.553685	•602219	3.744643
58	•667498	2.859681	•648059	3.122720	•631690	3.356611	•617579	3.568513	•605199	3.763157
59	•671369	2.857464	•651575	3.127019	•634494	3.366143	•620639	3.582411	•608103	3.780805
60	•675252	2.853913	•655090	3.130177	•638198	3.374657	•623388	3.595380	•610992	3.797587
61	•679149	2.849010	•658606	3.132182	•641442	3.382151	•626727	3.607418	•613869	3.813502
62	•683061	2.842729	•662125	3.133021	•644682	3.388616	•629756	3.618523	•616733	3.828554
63	•686993	2.835048	•665649	3.132683	•647918	3.394049	•632777	3.628692	•619585	3.842739
64	•690946	2.825941	•669180	3.131154	•651153	3.398439	•635791	3.637924	•622428	3.856062
65	•694923	2.815378	•672718	3.128417	•654387	3.401779	•638799	3.646212	•625261	3.868517
66	•698928	2.803330	•676267	3.124457	•657621	3.404062	•641802	3.653554	•628086	3.880105
67	•702962	2.789765	•679828	3.119256	•660858	3.405276	•644802	3.659943	•630903	3.890823
68	•707031	2.774646	•683402	3.112795	•664098	3.405411	•647798	3.665375	•633713	3.900670
69	•711136	2.757937	•686993	3.105052	•667343	3.404455	•650793	3.669843	•636518	3.909641
70	•715281	2.739597	•690601	3.096008	•670594	3.402397	•653788	3.673340	•639317	3.917734
71	•719471	2.719583	•694239	3.085633	•673853	3.399229	•656782	3.675858	•642113	3.924944
72	•723710	2.697448	•697880	3.073915	•677121	3.394916	•659779	3.677389	•644905	3.931269
73	•728001	2.674342	•701555	3.060815	•680399	3.389464	•662777	3.677926	•647695	3.936700
74	•732350	2.649011	•705257	3.046309	•683689	3.382851	•665780	3.677456	•650470	3.941234
75	•736762	2.621736	•708989	3.030366	•686993	3.375057	•668787	3.675971	•653271	3.944866
76	•741242	2.592636	•712753	3.012953	•690313	3.366065	•671800	3.673460	•656059	3.947588
77	•745797	2.561462	•716552	2.994036	•693648	3.355855	•674820	3.669910	•658849	3.950273
78	•750433	2.528200	•720389	2.973578	•697002	3.344406	•677848	3.665309	•661640	3.952723
79	•755157	2.492772	•724267	2.951539	•700376	3.331694	•680885	3.659643	•664433	3.950221
80	•759978	2.445091	•728189	2.927877	•703773	3.317698	•683933	3.652900	•667231	3.949228
81	•764905	2.415062	•732160	2.902548	•707194	3.302991	•686993	3.645062	•670033	3.947283
82	•769947	2.372584	•736183	2.875498	•710641	3.285747	•690065	3.636115	•672841	3.944380
83	•775116	2.327545	•740262	2.846681	•714116	3.267737	•693152	3.626041	•675654	3.940505
84	•780425	2.279829	•744403	2.816040	•717622	3.248332	•696255	3.614823	•678476	3.935548
85	•785889	2.229274	•748609	2.783514	•721161	3.227498	•699374	3.602441	•681305	3.922942
86	•791523	2.175756	•752886	2.749038	•724735	3.205202	•702512	3.588876	•684144	3.922942
87	•797347	2.119101	•757241	2.712957	•728347	3.181409	•705671	3.574106	•686993	3.915066
88	•803384	2.059123	•761679	2.673957	•732000	3.156078	•708850	3.558109	•689853	3.906158
89	•809659	1.995612	•766209	2.633195	•735697	3.129170	•712053	3.540862	•692726	3.896202
90	•816204	1.928333	•770337	2.599171	•739442	3.100642	•715536	3.522339	•695611	3.885182
91	•823055	1.8457020	•775572	2.544789	•743337	3.070442	•718536	3.502515	•698512	3.873063
92	•830258	1.741366	•780425	2.4966945	•747087	3.038526	•721820	3.481359	•701428	3.859887
93	•837469	1.701020	•785407	2.446626	•750995	3.008926	•725134	3.458844	•704361	3.845574
94	•845956	1.615568	•790530	2.393407	•754967	2.969323	•728482	3.434938	•707312	3.830134
95	•854608	1.524525	•795408	2.337340	•759006	2.931919	•731864	3.409608	•710283	3.815337
96	•863942	1.427314	•801258	2.278505	•763118	2.892559	•735284	3.382818	•713275	3.795765
97	•874112	1.323225	•806899	2.215605	•767310	2.851173	•738745	3.354532	•716289	3.776797
98	•885336	1.213393	•812753	2.150961	•771587	2.807626	•742246	3.324709	•719326	3.756637

99	•897925	1.090670	•814846	2.081965	•775956	2.762015	•745797	3.293308	•722389	3.735173
100	•912361	•959617	•825208	2.009181	•780425	2.714071	•749395	3.260283	•725479	3.712468
101	•929439	•816199	•831876	1.932342	•785003	2.663758	•753046	3.225588	•728598	3.688466
102	•950630	•657447	•838895	1.851142	•789700	2.610970	•766753	3.189171	•731747	3.663135
103	•979185	•478549	•846320	1.765228	•794526	2.555594	•760520	3.150978	•734929	3.636448
104	1.025248	•270157	•854219	1.674190	•799495	2.497503	•764352	3.110952	•738145	3.608370
105	1.068775	•148129	•862681	1.577545	•804619	2.436560	•768253	3.069029	•741398	3.578869
106			•871821	1.474716	•809915	2.372613	•772228	3.025145	•744691	3.547908
107			•881792	1.364999	•815403	2.305493	•776283	2.979226	•748025	3.515449
108			•892810	1.247527	•821103	2.235013	•780425	2.931197	•751403	3.481451
109			•905186	1.121192	•827042	2.160965	•784660	2.880974	•754829	3.445872
110			•919397	•984539	•833250	2.083110	•788996	2.828467	•758304	3.408666
111			•936234	•835571	•839763	2.001183	•793442	2.773580	•761834	3.369785
112			•957161	•671361	•846626	1.914878	•798007	2.716206	•765421	3.329178
113			•985410	•487178	•853893	1.823843	•802702	2.656231	•769069	3.286788
114			1.031086	•273860	•861632	1.727671	•807539	2.593528	•772783	3.242558
115			1.074342	•149661	•869932	1.625878	•812532	2.527959	•776566	3.196425
116					•878905	1.517892	•817698	2.459372	•780425	3.148322
117					•888706	1.403015	•823055	2.387597	•784365	3.098177
118					•899548	1.280385	•828625	2.312448	•788392	3.045911
119					•911739	1.148906	•834432	2.233715	•792513	2.991443
120					•925756	1.007138	•840508	2.151161	•796735	2.934680
121					•942384	•853109	•846888	2.064520	•801067	2.875526
122					•963080	•683932	•853616	1.973488	•805519	2.813874
123					•991064	•494952	•860747	1.877714	•810101	2.749608
124					1.036401	•277180	•868348	1.776791	•814826	2.682213
125					1.079418	•151029	•876506	1.670241	•819707	2.612719
126							•885336	1.557492	•824760	2.539805
127							•894988	1.437853	•830004	2.463690
128							•905676	1.310468	•835460	2.384187
129							•917706	1.174252	•841153	2.301088
130							•931553	1.027780	•847113	2.214157
131							•947998	•869104	•853377	2.123127
132							•968491	•695375	•859988	2.027694
133							•996240	•502011	•867000	1.927511
134							1.041277	•280183	•874480	1.822171
135							1.084083	•152260	•882515	1.711199
136									•891218	1.594028
137									•900740	1.469970
138									•911292	1.338177
139									•923180	1.197575
140									•936876	1.046753
141									•953158	•883785
142									•973472	•705861
143									1.001011	•508465
144									1.045780	•282917
145									1.088398	•153377

RESPONDING	NUMBER	N=155			N=165			N=175			N=185			N=195		
		BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	BURRIT	WEIGHT	
0		•212346	•292282	•209635	•292903	•207463	•314357	•202560	•321235							
1		•244892	•439261	•241759	•450741	•238848	•461833	•206132	•472544							
2		•282536	•659306	•278908	•676652	•275538	•693390	•272395	•709556							
3		•307285	•835162	•332005	•852278	•299648	•878602	•899213	•926095							
4		•326224	•988926	•322005	•1.013232	•318087	•1.038588	•314434	•311014							
5		•341776	•1.122687	•337339	•1.152808	•333220	•1.181832	•329381	•1.209868							
6		•355084	•1.246725	•350458	•1.280390	•346164	•1.312824	•342161	•1.344155							
7		•366787	•1.361628	•361989	•1.398639	•357538	•1.434292	•353391	•1.468711							
8		•372777	•1.469109	•372324	•1.509311	•367729	•1.548030	•363349	•1.585395							
9		•386818	•1.573087	•381719	•1.613649	•376992	•1.655306	•372589	•1.695506							
10		•395592	•1.666352	•390357	•1.712574	•385505	•1.757069	•380988	•1.800000							
11		•403734	•1.757687	•398370	•1.806783	•393401	•1.854029	•388775	•1.899604							
12		•411344	•1.844935	•405858	•1.896828	•400776	•1.946754	•402881	•2.086349							
13		•418501	•1.928519	•412897	•1.983147	•407708	•2.121213	•409335	•2.174326							
14		•425266	•2.008796	•419550	•2.066105	•414257	•2.212123	•415457	•2.259146							
15		•431690	•2.086052	•425864	•2.146001	•420471	•2.203622	•421288	•2.341077							
16		•437813	•2.160541	•431880	•2.222083	•426391	•2.283181	•421288	•2.341077							
17		•443668	•2.232466	•437631	•2.297568	•432048	•2.360108	•426859	•2.420341							
18		•449285	•2.302011	•443146	•2.369644	•437470	•2.434593	•432197	•2.497134							
19		•454686	•2.369330	•448448	•2.439465	•442682	•2.506801	•437327	•2.571618							
20		•459894	•2.434560	•453558	•2.507177	•447703	•2.576874	•442267	•2.643945							
21		•464924	•2.497822	•458492	•2.572904	•452550	•2.644940	•447035	•2.714242							
22		•469754	•2.559221	•463221	•2.636750	•457232	•2.711107	•451646	•2.782620							
23		•474516	•2.618853	•467894	•2.698815	•461789	•2.775478	•456112	•2.849192							
24		•479101	•2.676800	•472387	•2.759185	•466191	•2.838142	•460446	•2.914038							
25		•483562	•2.733141	•476755	•2.817937	•470477	•2.899183	•464657	•2.977247							
26		•487907	•2.784267	•481009	•2.875145	•474648	•2.958663	•468754	•3.039038							
27		•492144	•2.841267	•485155	•2.930873	•478713	•3.016659	•472746	•3.099038							
28		•496281	•2.893170	•489202	•2.985174	•482679	•3.073225	•476639	•3.157748							
29		•500325	•2.943704	•493156	•3.038106	•486553	•3.128413	•480440	•3.215077							
30		•504282	•2.992915	•497023	•3.093106	•490341	•3.182278	•484156	•3.271080							
31		•508158	•3.040845	•500809	•3.140044	•494047	•3.234866	•487791	•3.325800							
32		•511957	•3.087537	•504519	•3.189135	•497678	•3.286215	•491350	•3.379280							
33		•515685	•3.133023	•508158	•3.237029	•501237	•3.336366	•494838	•3.431561							
34		•519345	•3.177338	•511729	•3.283756	•504729	•3.385353	•498260	•3.482681							
35		•522942	•3.220513	•515237	•3.329352	•508158	•3.433212	•501618	•3.532671							
36		•526480	•3.262573	•518684	•3.373842	•511527	•3.479972	•504916	•3.581568							
37		•529961	•3.303551	•522076	•3.417256	•514839	•3.525662	•508158	•3.629396							
38		•533388	•3.343464	•525414	•3.459614	•518098	•3.571888	•511346	•3.676188							
39		•536766	•3.382339	•528701	•3.500954	•521305	•3.613935	•514833	•3.721918							
40		•540096	•3.420195	•531940	•3.541282	•524465	•3.656567	•517573	•3.766747							
41		•543381	•3.457050	•535134	•3.580626	•527616	•3.698223	•520616	•3.810566							
42		•546623	•3.492926	•538285	•3.619005	•530650	•3.738923	•523616	•3.853435							
43		•549825	•3.527836	•541395	•3.656435	•533680	•3.778686	•526574	•3.895380							
44		•552988	•3.561796	•544466	•3.692932	•536670	•3.817533	•529433	•3.936413							
45		•556115	•3.594821	•547500	•3.728513	•539623	•3.856475	•532375	•3.976554							
46		•559208	•3.626923	•550499	•3.763190	•542540	•3.892528	•535220	•4.015819							
47		•562268	•3.658115	•553464	•3.796978	•545423	•3.928711	•538031	•4.054224							
48		•565297	•3.688408	•556398	•3.829891	•548275	•3.964031	•540810	•4.091779							

49	•56R296	3.717814	•559301	3.861938	•551095	3.998506	•543557	4.128502	•536594	4.252724
50	•571267	3.746342	•562176	3.893130	•555885	4.033142	•566274	4.164402	•539246	4.290735
51	•574212	3.773999	•565022	3.923478	•556648	4.066955	•548963	4.199495	•541870	4.327498
52	•577131	3.800794	•567843	3.952691	•559379	4.096952	•551625	4.233785	•544466	4.363374
53	•580027	3.826735	•570639	3.981677	•562094	4.128145	•554260	4.267289	•547036	4.400023
54	•582900	3.851829	•573411	4.009544	•564780	4.158540	•556871	4.300013	•549580	4.434910
55	•585432	3.876080	•576161	4.036600	•567042	4.188149	•559497	4.331968	•552100	4.459039
56	•588582	3.899497	•578889	4.062550	•570082	4.216976	•554201	4.363158	•554596	4.492420
57	•591393	3.922083	•581596	4.088303	•572701	4.245029	•564563	4.393596	•557071	4.535066
58	•594187	3.943843	•584385	4.112963	•575299	4.272315	•567083	4.423290	•559523	4.569811
59	•596963	3.964779	•586954	4.136834	•577878	4.298842	•569584	4.452244	•561955	4.608649
60	•599723	3.984896	•589606	4.159923	•580438	4.324612	•572066	4.480463	•564368	4.648649
61	•602468	4.004198	•592242	4.182232	•582981	4.349635	•574529	4.507957	•566761	4.688417
62	•605199	4.022685	•594662	4.203767	•585507	4.373910	•576974	4.534729	•569137	4.687484
63	•607916	4.040361	•597466	4.224530	•588017	4.397445	•579403	4.560786	•571495	4.715855
64	•610620	4.057227	•600057	4.244523	•590512	4.420244	•581815	4.586130	•573836	4.743533
65	•613313	4.073282	•602634	4.263751	•592992	4.442310	•584212	4.610766	•576161	4.770527
66	•615995	4.088528	•605199	4.282213	•595458	4.463646	•586595	4.634703	•578470	4.796840
67	•618666	4.102966	•607751	4.299914	•597911	4.484256	•588963	4.657940	•580765	4.822475
68	•621329	4.116595	•610293	4.316853	•600352	4.504140	•591318	4.680480	•583046	4.847439
69	•623983	4.129412	•612824	4.333030	•602781	4.523300	•593660	4.702330	•585314	4.871734
70	•626629	4.141421	•615345	4.348448	•605199	4.541741	•595990	4.723488	•587568	4.895365
71	•629268	4.152617	•617858	4.363106	•607606	4.559463	•598308	4.743960	•589810	4.918334
72	•631901	4.162998	•620358	4.377004	•610003	4.576468	•600615	4.763745	•592040	4.940642
73	•634528	4.172563	•622858	4.390143	•612391	4.592755	•602912	4.782848	•594258	4.962297
74	•637150	4.181307	•625347	4.402520	•614770	4.608323	•605199	4.801269	•596466	4.983297
75	•639768	4.189230	•627829	4.414134	•617141	4.623178	•607476	4.819011	•598664	5.003644
76	•642383	4.196326	•630306	4.424984	•619504	4.637314	•609744	4.836074	•600851	5.023344
77	•644995	4.202592	•632777	4.435068	•621860	4.650735	•612004	4.852457	•603029	5.042395
78	•647605	4.208024	•635243	4.444386	•624210	4.663438	•614256	4.868162	•605199	5.066797
79	•650214	4.212617	•637706	4.452932	•626553	4.675423	•616501	4.883190	•607359	5.078557
80	•652822	4.216365	•640165	4.460703	•628891	4.686689	•618738	4.897540	•609512	5.095671
81	•655430	4.219264	•642621	4.467699	•631224	4.697233	•620970	4.911213	•611657	5.112141
82	•658039	4.221306	•645174	4.473916	•633553	4.707054	•623195	4.924207	•613795	5.127969
83	•660649	4.222485	•647526	4.479347	•635877	4.716151	•625414	4.936522	•615926	5.143152
84	•663261	4.222795	•649977	4.483991	•638198	4.724520	•627628	4.948157	•618051	5.157694
85	•665877	4.222828	•652427	4.488784	•640516	4.732160	•629838	4.959110	•620169	5.171593
86	•668496	4.220776	•654877	4.490896	•642831	4.739067	•632043	4.969381	•622282	5.184847
87	•671119	4.218431	•657327	4.493146	•645144	4.745239	•634244	4.978969	•624390	5.197459
88	•673788	4.215183	•659779	4.494587	•647456	4.750670	•636442	4.987868	•626493	5.209425
89	•676382	4.211026	•662232	4.495215	•649767	4.755359	•638637	4.996080	•628592	5.220745
90	•679023	4.205946	•664688	4.496920	•652077	4.759301	•640829	5.003602	•630686	5.231419
91	•681671	4.199934	•666908	4.498399	•654387	4.762491	•643079	5.010428	•632777	5.241444
92	•684327	4.185071	•669246	4.499214	•656697	4.764925	•645207	5.016561	•634864	5.250821
93	•686993	4.185071	•672074	4.485893	•659008	4.766506	•647579	5.022993	•636940	5.265545
94	•689668	4.176196	•674545	4.488593	•661321	4.767506	•649579	5.026722	•639030	5.276114
95	•692354	4.166342	•677021	4.481484	•663635	4.767641	•651765	5.030744	•641110	5.275029
96	•695052	4.155495	•679504	4.476207	•665952	4.766999	•653950	5.034057	•643187	5.281787
97	•697762	4.143641	•681992	4.470054	•668271	4.765572	•656135	5.036655	•645263	5.287882
98	•700486	4.130766	•684489	4.463013	•670594	4.763355	•658321	5.038535	•647337	5.293313

99	• 7032224	4.116854	• 686993	4.455075	• 672921	4.760341	• 660508	5.039690	• 649411	5.298079
100	• 7059777	4.101888	• 689506	4.446230	• 675252	4.556522	• 662696	5.040119	• 651484	5.302174
101	• 708747	4.083852	• 692028	4.436466	• 677588	4.751891	• 664887	5.039813	• 653557	5.305597
102	• 711535	4.068217	• 694560	4.425770	• 679930	4.474640	• 667080	5.038769	• 655631	5.308342
103	• 714342	4.050495	• 697104	4.414131	• 682277	4.470160	• 669275	5.036980	• 657704	5.310406
104	• 717168	4.031135	• 699659	4.401537	• 684631	4.733044	• 671474	5.034441	• 659779	5.311785
105	• 720015	4.010627	• 702226	4.387972	• 686993	4.725080	• 673676	5.031145	• 661854	5.312475
106	• 722886	3.988449	• 704807	4.373423	• 689362	4.716261	• 675883	5.027085	• 663932	5.312742
107	• 725740	3.966019	• 707402	4.357875	• 691739	4.706576	• 678094	5.022256	• 666011	5.311769
108	• 728699	3.944991	• 710012	4.341314	• 694126	4.696913	• 680310	5.016648	• 668093	5.310362
109	• 731645	3.916661	• 712638	4.323722	• 696521	4.684564	• 682531	5.010256	• 670177	5.308247
110	• 734619	3.890052	• 715281	4.305082	• 698928	4.672217	• 684759	5.003071	• 672264	5.305418
111	• 737624	3.862166	• 717942	4.285376	• 701345	4.658958	• 686993	4.995084	• 674355	5.301868
112	• 740650	3.832943	• 720622	4.264587	• 703773	4.644777	• 689234	4.986289	• 676449	5.297592
113	• 743731	3.802362	• 723322	4.242695	• 706214	4.629659	• 691482	4.976673	• 678548	5.292584
114	• 746837	3.770391	• 726044	4.219678	• 708668	4.613593	• 693738	4.966232	• 680651	5.286436
115	• 749941	3.736994	• 728788	4.195516	• 711136	4.596562	• 696003	4.954951	• 682760	5.280343
116	• 753165	3.702135	• 731555	4.170186	• 713618	4.578553	• 698276	4.942823	• 684873	5.273096
117	• 756392	3.665775	• 734348	4.143665	• 716116	4.559549	• 700559	4.929837	• 686993	5.265089
118	• 759664	3.627874	• 737167	4.115928	• 718630	4.539537	• 702853	4.915981	• 689118	5.256314
119	• 762985	3.586386	• 740013	4.086948	• 721161	4.518497	• 705157	4.901245	• 691251	5.246763
120	• 766356	3.547267	• 742890	4.056700	• 723710	4.496641	• 707473	4.885616	• 693390	5.236427
121	• 769782	3.504468	• 745797	4.025154	• 726278	4.473267	• 709800	4.869082	• 695537	5.225298
122	• 773267	3.459935	• 748737	3.992279	• 728866	4.449040	• 712140	4.851630	• 697692	5.213366
123	• 776813	3.413615	• 751712	3.958045	• 731476	4.423709	• 714494	4.833247	• 699856	5.200623
124	• 780425	3.365448	• 754724	3.922418	• 734107	4.397257	• 716861	4.813918	• 702028	5.187058
125	• 784108	3.315370	• 757774	3.885363	• 736762	4.369660	• 719244	4.793629	• 704210	5.172663
126	• 787867	3.263314	• 760866	3.846842	• 739442	4.340896	• 721642	4.772366	• 706402	5.157424
127	• 791708	3.209207	• 764001	3.806816	• 742147	4.310941	• 724056	4.750111	• 708605	5.141332
128	• 795635	3.152971	• 767182	3.765244	• 744880	4.279768	• 726487	4.726849	• 710819	5.124377
129	• 799657	3.094522	• 770412	3.722082	• 747641	4.247353	• 728937	4.702562	• 713044	5.106544
130	• 803781	3.033170	• 773693	3.677283	• 750433	4.213667	• 731405	4.677232	• 715281	5.087824
131	• 808015	2.970616	• 777030	3.630797	• 753256	4.178681	• 733893	4.650841	• 717532	5.068202
132	• 812369	2.904953	• 780425	3.582574	• 756114	4.142364	• 736402	4.623369	• 719796	5.047664
133	• 816853	2.836667	• 783883	3.532555	• 759006	4.104686	• 738933	4.594796	• 722074	5.026198
134	• 821479	2.765629	• 787407	3.480682	• 761936	4.065611	• 741487	4.565100	• 724367	5.003790
135	• 826262	2.691703	• 791003	3.426891	• 764905	4.025104	• 744065	4.534259	• 726675	4.980423
136	• 831216	2.614733	• 794675	3.371113	• 767916	3.983128	• 746668	4.502250	• 729000	4.959208
137	• 836361	2.534553	• 798428	3.313277	• 770970	3.939642	• 749298	4.469048	• 731341	4.904418
138	• 841717	2.450975	• 802269	3.253304	• 774072	3.894606	• 751955	4.434628	• 733701	4.870458
139	• 847311	2.363788	• 806205	3.191109	• 777222	3.847974	• 754642	4.398963	• 736079	4.847058
140	• 853170	2.272760	• 810242	3.126601	• 780425	3.799699	• 757360	4.362025	• 738477	4.848656
141	• 859332	2.177623	• 814390	3.059682	• 783684	3.749733	• 760110	4.323785	• 740895	4.819192
142	• 865840	2.078076	• 818658	2.990245	• 787001	3.698021	• 762894	4.284211	• 743335	4.788647
143	• 873772	1.973772	• 823055	2.918174	• 790381	3.644509	• 765714	4.243272	• 745797	4.757000
144	• 880119	1.864306	• 827595	2.843343	• 793829	3.589135	• 768572	4.200932	• 748283	4.724230
145	• 888045	1.749207	• 832291	2.765613	• 797347	3.531836	• 771470	4.157157	• 750793	4.690313
146	• 896635	1.627909	• 837158	2.684830	• 800942	3.464830	• 774410	4.111907	• 753329	4.655225
147	• 906041	1.499732	• 842216	2.600827	• 804619	3.411184	• 777394	4.065144	• 755893	4.618944
148	• 916472	1.363835	• 847484	2.513416	• 808383	3.347679	• 780425	4.0106825	• 758484	4.581442

149	•928234	1.219151	•852988	2.422388	•812243	3.281945	•783506	3.966905	•761105	4.542693
150	•941795	1.064286	•858758	2.327511	•816204	3.213389	•786640	3.915338	•763758	4.502668
151	•957432	•897335	•864829	2.228518	•820275	3.143414	•789829	3.862072	•766443	4.461336
152	•978083	•715523	•871245	2.125108	•824466	3.070412	•793077	3.807055	•769163	4.418669
153	1.005435	•514399	•878058	2.016936	•828187	2.994768	•796389	3.750231	•771920	4.374633
154	1.049963	•285422	•885336	1.903602	•833250	2.916354	•799767	3.691539	•774714	4.329191
155	1.092410	•154397	•893163	1.784633	•837669	2.835333	•803218	3.630915	•775548	4.282310
156			•901653	1.659471	•842659	2.750651	•806744	3.568290	•780425	4.233951
157			•910954	1.527438	•847638	2.663040	•810353	3.503591	•783347	4.184073
158			•921277	1.387702	•852828	2.572014	•814049	3.436740	•786316	4.132634
159			•932925	1.239204	•858252	2.477364	•817841	3.367650	•789334	4.079589
160			•946365	1.080566	•863842	2.378857	•821734	3.296233	•792406	4.024891
161			•962370	•909901	•869932	2.276229	•825737	3.222288	•795533	3.968489
162			•982376	•724471	•876264	2.169181	•829859	3.146009	•798720	3.910330
163			1.009557	•519883	•882993	2.057368	•834112	3.066979	•801971	3.850358
164			1.053866	•287730	•890185	1.940392	•838506	2.985173	•805288	3.788511
165			1.096158	•155332	•897625	1.817784	•843055	2.900451	•808677	3.724726
166					•906323	1.688988	•847775	2.812661	•812142	3.658934
167					•915531	1.553334	•852684	2.721636	•815690	3.591061
168					•925756	1.409994	•857803	2.627189	•819325	3.521029
169					•937300	1.257919	•863157	2.529112	•823055	3.448752
170					•950630	1.095745	•868774	2.427173	•826887	3.374139
171					•966516	•921606	•874690	2.321110	•830829	3.297093
172					•986389	•732793	•880948	2.210623	•834889	3.217504
173					1.013416	•524974	•887601	2.095371	•839880	3.135260
174					1.057525	•289865	•894715	1.974956	•843412	3.050231
175					•9099676	•156195	•902374	1.848914	•847899	2.962280
176							•910691	1.716692	•852556	2.871256
177							•919813	1.577625	•857401	2.776990
178							•929948	1.430890	•862456	2.679297
179							•941399	1.275450	•867745	2.577969
180							•954628	1.109952	•873296	2.472776
181							•970406	•932549	•879145	2.363456
182							•990157	•740564	•885336	2.249711
183							1.017042	•529719	•891919	2.131201
184							1.060967	•291850	•898962	2.007531
185							1.102989	•156994	•906549	1.878239
186									•914790	1.742777
187									•923834	1.600483
188									•933888	1.450542
189									•945252	1.291925
190									•958390	1.133292
191									•974067	•942815
192									•993707	•747846
193									1.020461	•534158
194									•293700	•157738
195									1.106119	

11-20	FK	>	punch with a decimal
1-10	C		
column #	variable		

1st parameter card, FORMAT (8F10.6)

b) Parameter cards

a) control cards (machine installation dependent).

a job deck:

The user must specify values for all of the above variables. In addition to these will be needed a set of control cards which direct the computer to compile, link edit, and execute the program. The control cards will depend on the users particular machine installation. The following breakdown will illustrate the order and format for submitting a job deck:

C - the parameter c in the Burr distribution.
 FK - the parameter k in the Burr distribution.
 K - the number of dose levels.
 IT - the number of iterations, IT=1.
 SS(I) - the number of subjects at the Ith dose.
 OP(I) - the proportion of subjects responding at the Ith dose level.
 X(I) - the value of the Ith dose level.

read by the program is defined as follows:

The program given in this appendix may be used to find estimates for the parameters of a quantal bio-assay, where the quantal response is assumed to be either the Burr or logistic distribution. Input which is

COMPUTER PROGRAM FOR BURRIT AND LOGIT ANALYSIS

2nd parameter card, FORMAT (8I10)

$\frac{\text{variable}}{K}$	$\frac{1-10}{11-20}$
> punch right justified IT=1	

c) Data cards, FORMAT (F18.6, 2F10.6)

Ith data card, K data cards are needed.

$\frac{\text{variable}}{SS(I)}$	$\frac{1-18}{19-28}$
> punch with decimal X(I)	29-38

The parameter and data cards would take the following form for

analysis of Data Set Two given in Chapter I.

1st parameter card

$\frac{\text{variable}}{4.874}$	$\frac{1-10}{11-20}$
---------------------------------	----------------------

2nd parameter card

$\frac{\text{variable}}{5}$	$\frac{10}{20}$
-----------------------------	-----------------

Data Cards

$\frac{\text{Col. 1-18}}{40.}$	$\frac{\text{Col. 19-28}}{.175}$	$\frac{\text{Col. 29-38}}{0.0}$
40.	.800	1.386
40.	.875	2.079
40.	.950	2.772

DIMENSION X(100), OP(100), OQ(100), EP(100), EQ(100)
1, U(100), GAMA(100), SS(100), OB(100), EB(100)

1 FORMAT(8F10.6)

2 FORMAT(1H0)

3 FORMAT(1H1)

4 FORMAT(1H1)

5 FORMAT(60X, 'ITERATION NO', 2X, I2)

12 FORMAT(65X, 'C=', 2X, F10.6, 2X, 'K=', 2X, F10.6)

13 FORMAT(F18.6, 2F10.6)

14 FORMAT(///57X, 'BURRIT ANALYSIS')

15 FORMAT(///58X, 'LOGIT ANALYSIS')

16 FORMAT(6E20.6)

READ 2, C, FK

21 READ 1, K, IT

READ 13, (SS(I), OP(I), X(I), I=1, K)

IN STATEMENT 20 SET X(I) EQUAL TO THE DESIRED TRANSFORMATION

OF DOSE AND REMOVE THE 'C' FROM COLUMN 6 OF STATEMENTS 18, 19 AND 20

18 DO 20 I=1, K

19 TIPS=X(I)

20 X(I)=ALOG(TIPS)

C1=1./C

C2=1./FK

IF(OP(I).EQ.0) OP(I)=1./2.*SS(I)

IF(OP(K).EQ.1) OP(K)=1.-1./2.*SS(K)

DO 30 I=1, K

OQ(I)=1.-OP(I)

EP(I)=OP(I)

40 I=1

BURRIT ANALYSIS ADJUSTED FOR BIAS AND VARIANCE

50 CALL ADJUST(K, C, FK, C1, C2, SS, U, GAMA, EP, EQ)

DO 100 I=1, K

OB(I)=(1./OQ(I))*C2-1.)**C1

CALL PARAM(X, OB, U, GAMA, A, B, K)

PRINT 4

PRINT 14

PRINT 3

PRINT 5, L

PRINT 3

PRINT 12, C, FK

DO 101 I=1, K

EP(I)=1.-1./((1.+A+B*X(I))*C)**FK

CALL RESID(X, OB, U, GAMA, EB, K, A, B, EP, OP, SS)

I=L+1

IF(L-IT) 50, 50, 102

102 I=1

SUBROUTINE ADJUST(K,C,FK,C1,C2,SS,U,GAMA,EP,EQ)
 DIMENSION SS(1),U(1),GAMA(1),EP(1),EQ(1)

FORTRAN IV G LEVEL 1, MOD 3 ADJUST

```

C
C
C
DO 103 I=1,K
  EP(I)=OP(I)
  EQ(I)=1.-OP(I)
  U(I)=0.0
  DO 104 I=1,K
    GAMA(I)=(OB(I)*EP(I)*OB(I)*EP(I))/(C*C*FK*FK*OQ(I)*OQ(I)*(1.-OQ(I)**2))
    CALL PARAM(X,OB,U,GAMA,A,B,K)
    PRINT 4
    PRINT 14
    PRINT 3
    PRINT 5,L
    PRINT 3
    PRINT 12,C,FK
    DO 106 I=1,K
      EP(I)=1.-1./((1.+A+B*X(I)**C)**FK)
      CALL RESID(X,OB,U,GAMA,EB,K,A,B,EP,OP,SS)
      I=L+1
    IF(L-IT) 104,104,107
  107 I=1
C
C
C
LOGIT ANALYSIS WITH WEIGHT=NPQ
DO 108 I=1,K
  EP(I)=OP(I)
  TEMPL=OP(I)/OQ(I)
  OB(I)=ALOG(TEMPL)
  EQ(I)=1.-OP(I)
  DO 109 I=1,K
    GAMA(I)=1./((SS(I)*EP(I)*EQ(I)))
    CALL PARAM(X,OB,U,GAMA,A,B,K)
    PRINT 4
    PRINT 15
    PRINT 3
    PRINT 5,L
    DO 111 I=1,K
      EP(I)=1./((1.+EXP(-(A+B*X(I))))))
      CALL RESID(X,OB,U,GAMA,EB,K,A,B,EP,OP,SS)
      I=L+1
    IF(L-IT) 109,109,112
  112 GO TO 21
END

```

16 FORMAT(6E20.6)

A=-C2

A1=A-1.

A2=A-2.

A3=A-3.

A4=A-4.

B=C1

B1=B-1.

B2=B-2.

B3=B-3.

A11=A*B

A21=A*B1

A22=A*A1*B

A31=A*B2

A32=(2.*B-1.)*A*A1

A33=B*A*A1*A2

A41=A*B3

A42=(3.*B1*A*A1*A2

A44=B*A*A1*A2*A3

DO 105 I=1,K

EQ(I)=1.-EP(I)

Q=EQ(I)

QA1=(Q**A)-1.

F0=QA1**B

F1=- (A11*Q**A1*F0)/QA1

F2=- (A21*Q**A1*F1+A22*Q**A2*F0)/QA1

F3=- (A31*Q**A1*F2+A32*Q**A2*F1-A33*Q**A3*F0)/QA1

F4=- (A41*Q**A1*F3+A42*Q**A2*F2-A43*Q**A3*F1+A44*Q**A4*F0)/QA1

U(I)=(EP(I)*EQ(I)*F2)/(2.*SS(I)+(EP(I)*EQ(I)*H(I)*F3)/(6

1.*SS(I)*SS(I)+(EP(I)*EQ(I)*H(I)*H(I)*SS(I)*SS(I)+(EP(I)*

2EQ(I)* (1.-6.*EP(I)*EQ(I))/(24.*SS(I)**3.))*F4

GAMA(I)=(EP(I)*EQ(I)*F1*F1)/SS(I)+(F2*F2*(2.*SS(I)*EP(I)*EQ

11)*EQ(I)* (1.-6.*EP(I)*EQ(I))/(4.*SS(I)**3.)))+(F1*F2

2*EP(I)*EQ(I)*EQ(I)*H(I))/(SS(I)**2.)

105 CONTINUE

RETURN

END

FORTRAN IV G LEVEL 1, MOD 3 PARAM

SUBROUTINE PARAM(X,OB,U,GAMA,A,B,K)

DIMENSION X(1),OB(1),U(1),GAMA(1)

16 FORMAT(6E20.6)

SX2G=0.

SBUG=0.

SXG=0.

SBUG=0.

SG=0.

DO 51 I=1,K

```

SUBROUTINE RESID(X,OB,U,GAMA,EB,K,A,B,EP,OP,SS)
DIMENSION X(1),OB(1),U(1),GAMA(1),EB(1),EP(1),OP(1),SS(1)
3 FORMAT (1H0
6 FORMAT (11X,'OBSERVED',12X,'BIAS',14X,'WEIGHT',11X,'ESTIMATED',1
11X,'RESIDUAL',11X,'SSE/CELL')
7 FORMAT (6F20.6)
8 FORMAT (13X,'ALPHA',15X,'BETA',16X,'TTS',17X,'SSR',17X,'SSE',14X,'R
1-SQUARED')
9 FORMAT (///13X,'DOSE',13X,'NO SUBJECTS',9X,'OBSERVED p',10X,'ESTIM
ATED p',11X,'RESIDUAL',9X,'CHI-SQUARE/CELL')
10 FORMAT (60X,'CHI-SQUARE')
11 FORMAT (45X,F20.6)
PRINT 3
PRINT 6
SSELT=0.
TSS=0.
DO 52 I=1,K
EB(I)=A+B*X(I)
RB=OB(I)-EB(I)
SSE=((RB-U(I))* (RB-U(I)))/GAMA(I)
SSELT=SSELT+SSE
TSS=TSS+((OB(I)-U(I))* (OB(I)-U(I)))/GAMA(I)
GAMA(I)=1./GAMA(I)
52 PRINT 7, OB(I),U(I),GAMA(I),EB(I),RB,SSE
SSR=TSS-SSELT
R2=SSR/TSS
PRINT 3
PRINT 8
PRINT 7, A,B,TSS,SSR,SSELT,R2
PRINT 9
CHI2T=0.
DO 55 I=1,K
CHI2=OP(I)-EP(I)
CHI2T=CHI2T+CHI2
55 PRINT 7, X(I),SS(I),OP(I),EP(I),RP,CHI2
PRINT 3
PRINT 10
PRINT 11, CHI2T
RETURN
END
FORTRAN IV G LEVEL 1, MOD 3 RESID
51 SG=SG+1./GAMA(I)
SBUXG=SBUXG+(OB(I)-U(I))*X(I))/GAMA(I)
SXG=SXG+X(I)/GAMA(I)
SBUG=SBUG+(OB(I)-U(I))/GAMA(I)
SX2G=SX2G+(X(I)*X(I))/GAMA(I)
B=(SG*SBUXG-SXG*SBUG)/DIV
A=(SX2G*SBUG-SXG*SBUXG)/DIV
DIV=SG*SX2G-(SXG*SXG)

```

- [1] Berkson, J. "Application of the Logistic Function to Bio-Assay." Journal of the American Statistical Association, 39, 357-64, 1944.
- [2] Berkson, J. "A Statistically Precise and Relatively Simple Method of Estimating the Bio-Assay with Quantal Response, Based on the Logistic Function." Journal of the American Statistical Association, 48, 565-99, 1953.
- [3] Berkson, J. "Maximum Likelihood and Minimum Chi-Square Estimates of the Logistic Function." Journal of the American Statistical Association, 50, 130-62, 1955 A.
- [4] Berkson, J. "Estimate of the Integrated Normal Curve by Minimum Normal Chi-Square with Particular Reference to Bio-Assay." Journal of the American Statistical Association, 50, 529-49, 1955 B.
- [5] Berkson, J. "Tables for Use in Estimating the Normal Distribution Function by Normal Analysis." Biometrika, 44, 411-35, 1957.
- [6] Burr, I. W. "Cumulative Frequency Functions." Annals of Mathematical Statistics, 13, 215-32, 1942.
- [7] Burr, I. W. "A Useful Approximation to the Normal Distribution Function, with Application to Simulation." Technometrics, 9, 647-51, 1967.
- [8] Burr, I. W. and Cislak, P. J. "On a General System of Distributions--I. Its Curve Shape Characteristics. II. The Sample Median." The Journal of the American Statistical Association, 63, 627-35, 1968.
- [9] Finney, D. J. Probit Analysis. Second Edition, London: Cambridge at the University Press, 1952.
- [10] Finney, D. J. Statistical Method in Biological Assay. Second Edition, London: Charles Griffin and Company Limited, 1964.
- [11] Gart, J. J. and Zweifel, J. R. "On the Bias of Various Estimators of the Logit and its Variance with Application to Quantal Bio-Assay." Biometrika, 54, 181-87, 1967.
- [12] Garwood, F. "The Application of Maximum Likelihood to Dosage-Mortality Curves." Biometrika, 32, 46, 1941.

LIST OF REFERENCES

- [13] Hartley, H. O. "Exact Confidence Regions for the Parameters in Non-linear Regression Laws." Biometrika, 51, 347-53, 1964.
- [14] Hewlett, P. S. and Plackett, R. L. "Statistical Aspects of the Independent Joint Action of Poisons, Particularly Insecticides." II. Examination of the Data for Agreement with the Hypothesis. Annals of Applied Biology, 37, 527-52, 1950.
- [15] Kendall, M. and Stuart, A. The Advanced Theory of Statistics, Vol. 1. Second Edition, London: Charles Griffin and Company Limited, 1963.
- [16] Moore, R. H. and Zeigler, R. K. "The Use of Non-linear Regression Methods for Analyzing Sensitivity and Quantal Response Data." Biometrics, 23, 563, 1967.
- [17] Murray, C. A. "Dosage-Mortality in the Peet-Grady Method." Soap, 99, 1938.
- [18] Nelder, J. A. "Weighted Regression, Quantal Response Data, and Inverse Polynomials." Biometrics, 24, 979-85, 1968.
- [19] Zeigler, R. K. and Moore, R. H. "Multivariate Quantal Response Analysis Using Regression Methods." Presented at the 126th Annual Meeting of the American Statistical Association, 1966.

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13. ABSTRACT

If different levels of stimuli are applied and then counts are made of the number reacting and not-reacting to the stimuli, a sensitivity analysis is usually the best analysis to be applied. Sensitivity analysis may be used to determine the probability of destroying a target with different size bombs, without the necessity of actually having data on all the different sizes. It may also be used in determining the effect of a drug on a patient or the effect of pollution on a body of water. This paper is concerned with a generalized model for sensitivity analysis based on the Burr distribution. This generalization gives a model that will fit quite well data that comes from probit or logit models, but also covers many additional situations that the probit or logit models would not fit. A FORTRAN IV program is included for doing the analysis on a computer. Confidence intervals for any quantile are also given.